

## Session 13

Recall...

13.1

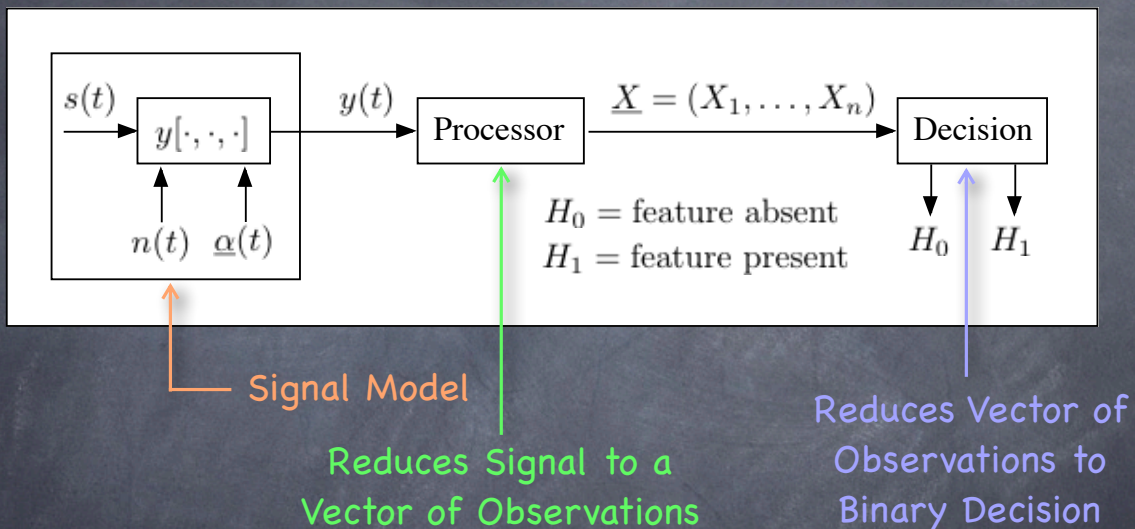
# Detection Theory

- 👁 The most basic task of a RADAR system is to **detect** whether or not an object (target) is present.
- 👁 We may also wish to detect if some "feature" is present in the scattered signal:
  - > Spectral Component
  - > signal family characteristic
  - > depolarized wave component -

## An Optimal Test?

- > How do we design an optimal test in this case?
- > We want as small a number of "false alarms" as possible.
- > We want as small a number of "misses" as possible.
- > Clearly these two goals contradict each other.
- > An "optimal" test must trade these two goals off against each other. But how?

## General Framework for Radar Detection



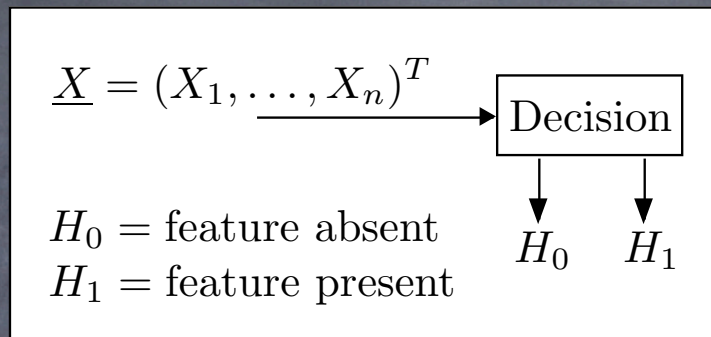
Input to processor is a random process  $y(t)$ .

Processor output  $\underline{X} = (X_1, \dots, X_N)^T$  is a vector of jointly distributed random variables—a *random vector*.



For now, we will focus on the decision block

13.4



We will assume that after processing by the processor, the random vector  $\underline{X}$  is described by a cdf parameterized a parameter vector  $\underline{\theta}$ :

$$F_{\underline{\theta}}(\underline{x}) = P_{\underline{\theta}}(\{X_1 \leq x_1\} \cap \dots \cap \{X_N \leq x_N\})$$

$\uparrow (x_1, x_2, \dots, x_n)^T$

How do we design optimal decision blocks?

## Binary Hypothesis Testing

13.5

Let

$$\underline{X} = (X_1, \dots, X_N)^T = \text{vector of observations}$$

governed by

$$F_{\underline{\theta}}(\underline{x}) = P_{\underline{\theta}}(\{X_1 \leq x_1\} \cap \dots \cap \{X_N \leq x_N\}),$$

where

$$P_{\underline{\theta}} = \text{probability measure parameterized by } \underline{\theta}$$



For example:

Suppose  $X_1, \dots, X_N$  are i.i.d Gaussians with mean  $\mu$  and variance  $\sigma^2$ .

$$\begin{aligned} F_{\underline{\theta}}(\underline{x}) &= F_{\underline{\theta}}(x_1, \dots, x_N) \\ &= F_{\underline{\theta}}(x_1) F_{\underline{\theta}}(x_2) \cdots F_{\underline{\theta}}(x_N) \\ &= \Phi\left(\frac{x_1 - \mu}{\sigma}\right) \Phi\left(\frac{x_2 - \mu}{\sigma}\right) \cdots \Phi\left(\frac{x_n - \mu}{\sigma}\right) \end{aligned}$$

where

$$\Phi(z) = \int_{-\infty}^z \frac{1}{\sqrt{2\pi}} e^{-v^2/2} dv,$$

and

$$\underline{\theta} = (\mu, \sigma^2)^T \in \Theta = \mathbf{R} \times [0, \infty)$$

Let  $\Theta_0$  and  $\Theta_1$  be a partition of  $\Theta$  (the parameter space) (i.e.,  $\Theta_0 \cap \Theta_1 = \emptyset$  and  $\Theta_0 \cup \Theta_1 = \Theta$ )

Either

$$\underline{\theta} \in \Theta_0 \quad \Rightarrow \quad H_0 \text{ is true}$$

or

$$\underline{\theta} \in \Theta_1 \quad \Rightarrow \quad H_1 \text{ is true}$$

We will assume both  $H_0$  and  $H_1$  are simple hypotheses.

**Defn.** A hypothesis  $H_i$  is said to be a *simple hypothesis* if its corresponding parameter set  $\Theta_i$  contains only a single element  $\underline{\theta}_i$ . If  $\Theta_i$  contains more than one element, then  $H_i$  is a *composite hypothesis*.



If  $H_0$  and  $H_1$  are both simple, then

$$\Theta_0 = \{\underline{\theta}_0\} \text{ and } \Theta_1 = \{\underline{\theta}_1\}$$

and

$$\Theta = \Theta_0 \cup \Theta_1 = \{\underline{\theta}_0, \underline{\theta}_1\}$$

Then if hypothesis  $H_i$  is true,  $\underline{X}$  has a single cdf  $F_{\underline{\theta}_i}(\underline{x})$ , for  $i = 0, 1$ .

For a composite  $H_i$ , there is a corresponding set of cdfs:

$$\{F_{\underline{\theta}}(\underline{x}); \underline{\theta} \in \Theta_i\}$$

Hypothesis  $H_i$  is true if any one of these cdfs describes  $\underline{X}$ .

Given a random vector  $\underline{X}$  having cdf  $F_{\underline{\theta}}(\underline{x})$ ,  $\underline{\theta} \in \Theta$ , we can define a binary test of the form

$$\phi(\underline{x}) = \begin{cases} 1, & \text{for } \underline{x} \in \mathcal{R}, \\ 0, & \text{for } \underline{x} \in \mathcal{A}, \end{cases}$$

where

$\mathcal{R} = \text{rejection region of } H_0$

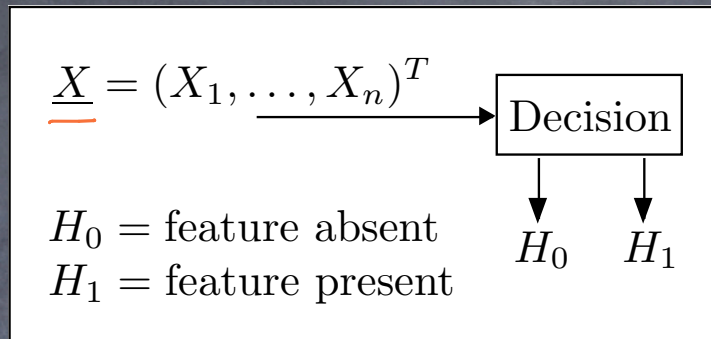
$\mathcal{A} = \text{acceptance region of } H_0$

Assuming real RVs  $X_1, \dots, X_N$ ,  $\underline{x} \in \mathbf{R}^N$ , and

$$\begin{array}{ll} \mathcal{R} \subset \mathbf{R}^N & \mathcal{R} \cap \mathcal{A} = \emptyset \\ \mathcal{A} \subset \mathbf{R}^N & \mathcal{R} \cup \mathcal{A} = \mathbf{R}^N \end{array}$$



## Recap: Binary Hypothesis tests



$\underline{X} = (X_1, \dots, X_N)^T = \text{vector of observations}$   
 governed by

$$F_{\underline{\theta}}(\underline{x}) = P_{\underline{\theta}}(\{X_1 \leq x_1\} \cap \dots \cap \{X_N \leq x_N\}),$$

We will assume both  $H_0$  and  $H_1$  are simple hypotheses.

$$\text{Under } H_0, \Theta_0 = \{\underline{\theta}_0\} \Rightarrow \underline{X} \sim F_{\underline{\theta}_0}(\underline{x})$$

$$\text{Under } H_1, \Theta_1 = \{\underline{\theta}_1\} \Rightarrow \underline{X} \sim F_{\underline{\theta}_1}(\underline{x})$$

$$\phi(\underline{x}) = \begin{cases} 1, & \text{for } \underline{x} \in \mathcal{R}, \\ 0, & \text{for } \underline{x} \in \mathcal{A}, \end{cases}$$

$\mathcal{R} = \text{rejection region of } H_0$

$\mathcal{A} = \text{acceptance region of } H_0$

Note the language:

If  $\phi(\underline{X}) = 0$ , we accept hypothesis  $H_0$ .

If  $\phi(\underline{X}) = 1$ , we reject hypothesis  $H_0$ .

~~(and effectively accept  $H_1$ )~~



In general, two types of errors can occur:

$H_0$  is true, but  $\phi(\underline{X}) = 1$ .      *Type I Error* or  
*False Alarm*

$H_1$  is true, but  $\phi(\underline{X}) = 0$ .      *Type II Error* or *Miss*

The probability of Type I error is called the size or probability of false alarm of the test.

It is given by

$$\alpha = P_{\underline{\theta}_0}(\{\phi(\underline{X}) = 1\})$$

when  $H_0$  is simple.

Note that the size can also be written as

$$\alpha = P_{\underline{\theta}_0}(\{\phi(\underline{X}) = 1\})$$

$$= \boxed{E_{\underline{\theta}_0}[\phi(\underline{X})]}$$

*n.b.* the subscripts  $\underline{\theta}_0$  mean these quantities are computed using cdf  $F_{\underline{\theta}_0}(\underline{x})$  or corresponding pdf  $f_{\underline{\theta}_0}(\underline{x})$

If  $H_0$  is composite, the size is defined as

$$\alpha = \sup_{\underline{\theta} \in \Theta_0} \{E_{\underline{\theta}}[\phi(\underline{X})]\}$$

Here “sup” means *supremum* (i.e., the least upper bound)



Note that

$$\alpha = \sup_{\underline{\theta} \in \Theta_0} \{E_{\underline{\theta}}[\phi(\underline{X})]\}$$

represents the worst-case false alarm for any possible  $\underline{\theta}_0 \in \Theta_0$ .

If  $H_1$  is in effect and  $\phi(\underline{X}) = 1 \Rightarrow$  Correct Decision

If  $H_1$  is simple, the probability of correctly deciding  $H_1$  is called the power or probability of detection of  $\phi(\cdot)$ :

$$\beta = P_{\underline{\theta}_1}(\{\phi(\underline{X}) = 1\}) = E_{\underline{\theta}_1}[\phi(\underline{X})]$$

If  $H_1$  is composite, we compute  $\beta$  for each possible  $\underline{\theta}_1 \in \Theta_1$ :

$$\beta(\underline{\theta}) = P_{\underline{\theta}}(\{\phi(\underline{X}) = 1\}) = E_{\underline{\theta}}[\phi(\underline{X})], \quad \forall \underline{\theta} \in \Theta_1$$

Note the asymmetry between  $H_0$  and  $H_1$ :

$$\alpha = \sup_{\underline{\theta} \in \Theta_0} \{E_{\underline{\theta}}[\phi(\underline{X})]\} \quad (\text{worst case})$$

$$\beta(\underline{\theta}) = E_{\underline{\theta}}[\phi(\underline{X})], \quad \forall \underline{\theta} \in \Theta_1 \quad (\text{each case})$$

Why? Fundamental nature of frequentist hypothesis test.



## A Radar Interpretation

$$\alpha = \sup_{\underline{\theta} \in \Theta_0} \{E_{\underline{\theta}}[\phi(\underline{X})]\} \quad (\text{worst case})$$

$\Rightarrow \alpha$  is maximum possible probability of false alarm.

$$\beta(\underline{\theta}) = E_{\underline{\theta}}[\phi(\underline{X})], \quad \forall \underline{\theta} \in \Theta_1 \quad (\text{each case})$$

$\Rightarrow$  Evaluate the probability of detection for each possible target.

Because  $\beta$  is the “probability of detection,” we sometimes write it as

$$P_D = \beta = \text{Probability of Detection}$$

The probability of deciding that  $H_0$  is in effect when in fact  $H_1$  is—a type II error—is called the *miss probability*

$$P_M = 1 - \beta = 1 - P_D$$

Because  $\alpha$  is the probability of false alarm,” we sometimes write it as

$$P_{FA} = \alpha = \text{Probability of False Alarm}$$