

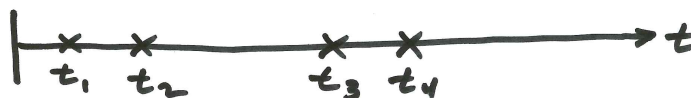
## Session 30

### The Poisson Point Process

30.1

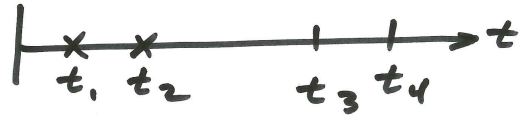
Defn: A point process is a set of random points  $\{t_i\}$  on the time axis.

n.b. The "points" are RVs that represent times at which some random "events" in time occur.  
(e.g. times at which light bulbs in your house burn out.)



A point process is a set of random points along the time axis:  $\{t_i\}$

30.2



$$t_1 \leq t_2 \leq t_3 \leq \dots \leq t_n \leq \dots$$

The points are ordered in time.

The  $\{t_i\}$  are a collection of RVs defined on some probability space  $(\mathcal{S}, \mathcal{F}, \mathcal{P})$ .

Defn: To each point process  $\{t_i\}$  we can assign a random process  $X(t)$  called the counting process, defined as

30.3

$$X(t) \triangleq \text{number of points in the interval } [0, t).$$

Defn: To any point process  $\{t_i\}$  we can associate a renewal process, a sequence of RVs defined by

30.4

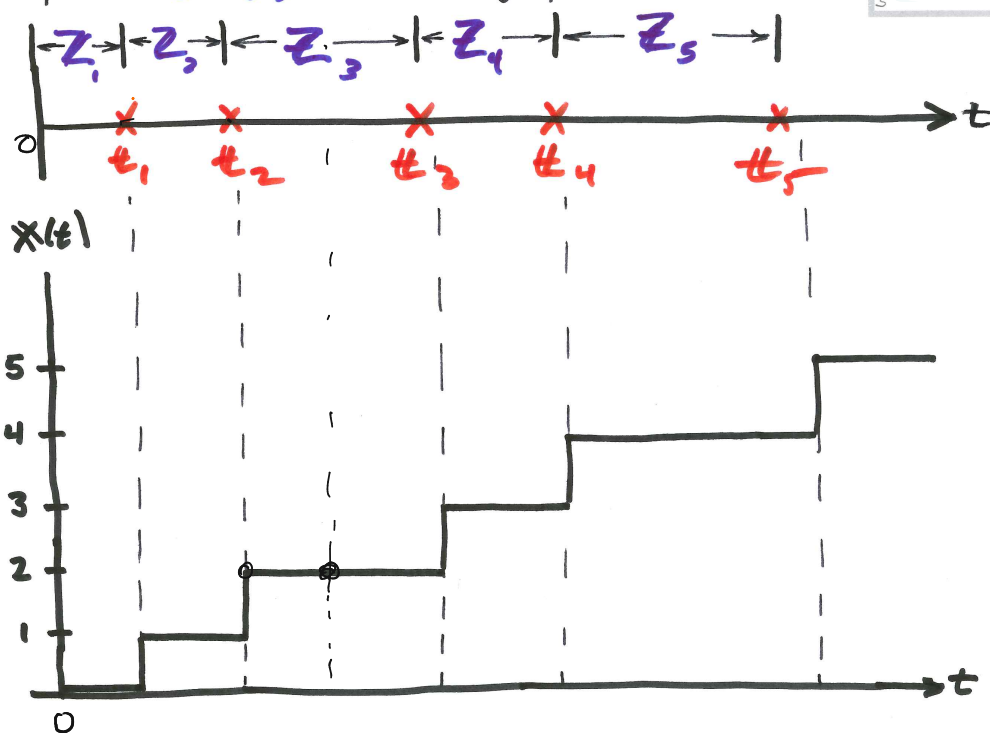
$$Z_n = \begin{cases} t_1, & n=1 \\ t_n - t_{n-1}, & n=2, 3, \dots \end{cases}$$

The renewal process represents the time between the events in the point process  $\{t_i\}$ .

The  $\{Z_n\}$  are RVs defined on  $(\mathcal{S}, \mathcal{F}, P)$ .

Relationship between point process  $\{t_i\}$ , renewal process  $\{Z_i\}$  and counting process  $X(t)$ :

30.5



Defn: A point process  $\{t_i\}$  is called a Poisson point process or set of Poisson points if

1. The number of points  $N(t_1, t_2)$  in the time interval  $[t_1, t_2)$  is a Poisson RV with mean  $\lambda(t_2 - t_1)$ , where

$$\lambda > 0: P(\{N(t_1, t_2) = k\}) = \frac{e^{-\lambda(t_2 - t_1)} [\lambda(t_2 - t_1)]^k}{k!}$$

for  $k = 0, 1, 2, \dots$

2. If  $[t_1, t_2) \cap [t_3, t_4) = \emptyset$ , then  $N(t_1, t_2)$  and  $N(t_3, t_4)$  are statistically independent RVs. (for real numbers  $t_1, t_2, t_3, t_4$ )

n.b. The mean is proportional to the length of the interval (i.e., the mean of  $N(t_1, t_2)$ ).

\* Technically, this is a homogeneous Poisson process with constant rate  $\lambda$ .

(This can be generalized to a non homogeneous Poisson process where the rate  $\lambda(t)$  is not constant. Here  $N(t_1, t_2)$  has mean  $\Lambda(t_1, t_2) = \int_{t_1}^{t_2} \lambda(t) dt$  we will not consider this)

Defn: The random process

$$X(t) \triangleq IN(0, t)$$

corresponding to a set of Poisson points  $\{t_i\}$  is called a Poisson Counting process

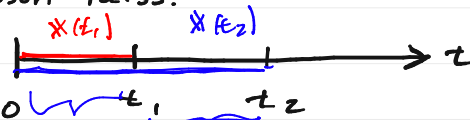
See Ex. 9-5 (10-5 in 3rd. Edition) of Papoulis.

For this R.P.  $X(t)$ , it can be shown that

$$E[X(t)] = \lambda t$$

$$R_{XX}(t_1, t_2) = \lambda \cdot \min(t_1, t_2) + \lambda^2 t_1 t_2.$$

n.b. In computing  $R_{XX}(t_1, t_2) = E[X(t_1)X(t_2)]$  we will use the independent increment property (prop. 2) of the Poisson process.

Assume  $t_2 > t_1$ : 

$$E[X(t_1)X(t_2)] = E[IN(0, t_1) \cdot IN(0, t_2)]$$

$$= E[X(t_1)(X(t_1) + (X(t_2) - X(t_1)))]$$

$$= E[X^2(t_1)] + E[X(t_1)(X(t_2) - X(t_1))]$$

$$= E[X^2(t_1)] + E[IN(0, t_1) \cdot IN(t_1, t_2)]$$

$$= E[X^2(t_1)] + E[IN(0, t_1)] \cdot E[IN(t_1, t_2)]$$

$$= \dots = \lambda t_1 + \lambda^2 t_1 t_2$$

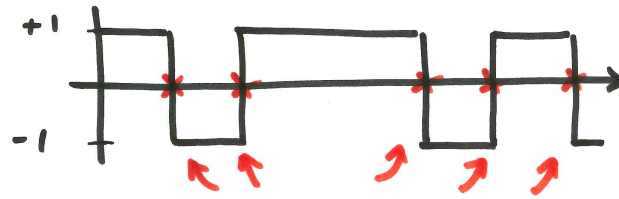
By symmetry, if  $t_1 > t_2$

$$E[X(t_1)X(t_2)] = \lambda t_2 + \lambda^2 t_1 t_2.$$

n.b. The Poisson process is used to generate the Random Telegraph Process

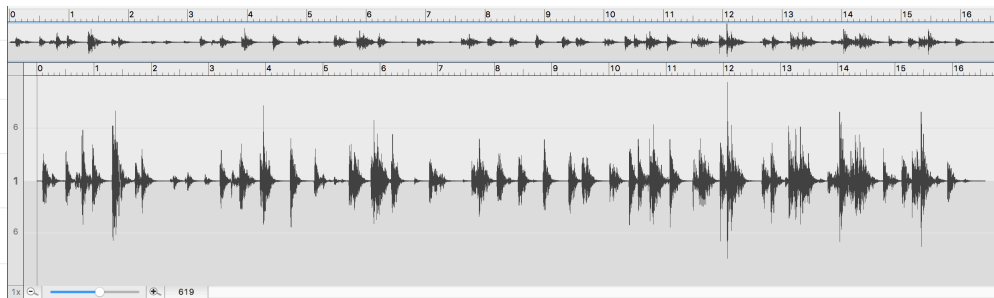
30.10  
10

eg.:



Poisson Points  
determine  
transitions

## Microwave Popcorn Recording



Q: What is the pdf of the  $n$ -th point  $t_n$  of a Poisson process?

CDF:  $F_{t_n}(t) = P(\{ \sum t_n \leq t \})$ ,  $t > 0$ .

Note that

$$\{ t_n \leq t \} \iff \{ N(0, t) \geq n \}$$

$$\Rightarrow \{ t_n \leq t \} = \{ N(0, t) \geq n \}$$

$$\begin{aligned} F_{t_n}(t) &\stackrel{\Delta}{=} P(\{ \sum t_n \leq t \}) = P(\{ N(0, t) \geq n \}) && 30.12 \\ &= 1 - P(\{ N(0, t) < n \}) \\ &= 1 - P(\{ N(0, t) \leq n-1 \}) \\ &= 1 - \sum_{k=0}^{n-1} \frac{e^{-\lambda t} (\lambda t)^k}{k!} \end{aligned}$$

Thus, the pdf of  $t_n$  is

$$\begin{aligned} f_{t_n}(t) &= \frac{d}{dt} (F_{t_n}(t)) \\ &= \frac{d}{dt} \left( 1 - \sum_{k=0}^{n-1} \frac{e^{-\lambda t} (\lambda t)^k}{k!} \right) \\ &= \frac{d}{dt} \left( \sum_{k=n}^{\infty} \frac{e^{-\lambda t} (\lambda t)^k}{k!} \right) \end{aligned}$$

$$\dots \sum_{k=n}^{\infty} \frac{d}{dt} \left( \frac{e^{-\lambda t} (\lambda t)^k}{k!} \right) \dots \quad (*)$$

30.13

Now

$$\frac{d}{dt} \left( \frac{e^{-\lambda t} (\lambda t)^k}{k!} \right) = -\frac{\lambda e^{-\lambda t} (\lambda t)^k}{k!} + \frac{k e^{-\lambda t} (\lambda t)^{k-1} \lambda}{k!}$$

$$= \lambda e^{-\lambda t} \left[ \frac{(\lambda t)^{k-1}}{(k-1)!} - \frac{(\lambda t)^k}{k!} \right] \leftarrow$$

So, substituting this expression back into  $(*)$ , we get

$$f_{t_n}(t) = \lambda e^{-\lambda t} \sum_{k=n}^{\infty} \left( \frac{(\lambda t)^{k-1}}{(k-1)!} - \frac{(\lambda t)^k}{k!} \right)$$

30.14

$$= \lambda e^{-\lambda t} \left( \underbrace{\frac{(\lambda t)^{n-1}}{(n-1)!} - \frac{(\lambda t)^n}{n!}}_{k=n} + \underbrace{\frac{(\lambda t)^n}{n!} - \frac{(\lambda t)^{n+1}}{(n+1)!}}_{k=n+1} + \underbrace{\frac{(\lambda t)^{n+1}}{(n+1)!} - \frac{(\lambda t)^{n+2}}{(n+2)!}}_{k=n+2} + \dots \right)$$

$$\lambda e^{-\lambda t} \left( \frac{(\lambda t)^{n-1}}{(n-1)!} \right)$$

$$\therefore f_{t_n}(t) = \frac{\lambda (\lambda t)^{n-1} e^{-\lambda t}}{(n-1)!}, \quad t \geq 0.$$



∴ The  $n$ -th point  $t_n$  of a homogeneous Poisson process with rate  $\lambda$  is

30.15

$$f_{t_n}(t) = \frac{\lambda^n t^{n-1} e^{-\lambda t}}{(n-1)!} \cdot \mathbb{1}_{[0, \infty)}(t)$$

This is the  $n$ -th order Erlang pdf  
(Agner Krarup Erlang - Danish mathematician and engineer)

It is a special case of the gamma pdf:

$$f_{\#}(x) = \frac{\beta^\alpha x^{\alpha-1} e^{-\beta x}}{\Gamma(\alpha)} \mathbb{1}_{[0, \infty)}(x) \quad \left( \phi_{\#}(s) = \frac{\beta^\alpha}{(s+\beta)^\alpha} \right)$$

where  $\alpha = n$  and  $\beta = \lambda$ .

(i.b.  $\Gamma(n) = (n-1)!$  for  $n \in \mathbb{N}$ )

For  $n=1$ , this becomes

30.16

$$f_{t_1}(t) = \lambda e^{-\lambda t} \cdot \mathbb{1}_{[0, \infty)}(t)$$

⇒  $t_1$  is exponentially distributed with mean  $1/\lambda$ .

⇒  $Z_n = t_n - t_{n-1}$  is exponentially distributed with mean  $1/\lambda$  when  $n = 2, 3, \dots$

(i.e., the time between Poisson Points is exponentially distributed with mean  $1/\lambda$ .)

## Simulating a Homogeneous Poisson Process

30.17

We can use the fact that the time between Poisson points in a homogeneous Poisson process with rate  $\lambda$  are independent identically distributed exponentially distributed RVs with mean  $1/\lambda$  to simulate a homogeneous Poisson Process:

1. Generate i.i.d. exponential RVs with mean  $\mu = 1/\lambda$ :  $Z_1, Z_2, \dots, Z_n$
2. Generate the Poisson points

$$t_k = \sum_{m=1}^k Z_m, \quad k=1, \dots, n$$

30.18

2. Generate the Poisson points

$$t_k = \sum_{m=1}^k Z_m, \quad k=1, \dots, n$$

Then we have that  $\{t_1, t_2, \dots, t_n\}$  are the first  $n$  Poisson points of a Poisson point process.

The i.i.d. exponential RVs with mean  $1/\lambda$

$$Z_1, Z_2, \dots, Z_n, \dots$$

are the renewal process corresponding to the Poisson points  $t_1, t_2, \dots, t_n, \dots$

## The End!

\* Please fill out the course evaluations for the course.

n.b. Overflow Section will take the final exam with FNY Section in G140

30.19

If you are looking for additional problems to solve for practice for the final exam, here is one source:

Hwei P. Hsu, Probability, Random Variables, and Random Processes, 3rd Edition, Schaum's Outline, McGraw-Hill, 2014

An electronic (Kindle) version is available for instant download on Amazon (\$14.16)

406 Fully solved problems.

## ECE 600 Final Exam

Monday, April 29, 2024

3:30 to 5:30 pm

Room FNY G140 { For FNY and  
ONC (overflow) Sections

This will be a comprehensive exam.

This is a closed book, closed notes exam.

You may not use a calculator.

Start each problem on a new page.

- 10 "simple" multiple choice\* questions 50%  
\* There will be room to show your work.
- Two "work-out" problems 34% ←
- One page of True/False Questions 16%

# EPE Online Section Answer Form

ECE600 Random Variables and Waveforms  
Spring ~~2022~~ **2024**

Mark R. Bell  
MSEE 336

## Final Exam Online Section Answer Form

~~May 4, 2022~~  
**April 29, 2024**

1-10	
11	
12	
13	
<b>Total</b>	

$$R(\tau) = \frac{1}{2\pi} \int_{-\infty}^{\infty} S(\omega) e^{j\omega\tau} d\omega \leftrightarrow S(\omega) = \int_{-\infty}^{\infty} R(\tau) e^{-j\omega\tau} d\tau$$

$\delta(\tau) \leftrightarrow 1$ $e^{j\beta\tau} \leftrightarrow 2\pi\delta(\omega - \beta)$ $e^{-\alpha \tau } \leftrightarrow \frac{2\alpha}{\alpha^2 + \omega^2}$ $e^{-\alpha \tau } \cos \beta\tau \leftrightarrow \frac{\alpha}{\alpha^2 + (\omega - \beta)^2} + \frac{\alpha}{\alpha^2 + (\omega + \beta)^2}$ $2e^{-\alpha\tau^2} \cos \beta\tau \leftrightarrow \sqrt{\frac{\pi}{\alpha}} [e^{-(\omega - \beta)^2/4\alpha} + e^{-(\omega + \beta)^2/4\alpha}]$ $\begin{cases} 1 - \frac{ r }{T} &  r  < T \\ 0 &  r  > T \end{cases} \leftrightarrow \frac{4 \sin^2(\omega T/2)}{T\omega^2}$ $\frac{\sin \sigma\tau}{\pi\tau} \leftrightarrow \begin{cases} 1 &  \omega  < \sigma \\ 0 &  \omega  > \sigma \end{cases}$	$1 \leftrightarrow 2\pi\delta(\omega)$ $\cos \beta\tau \leftrightarrow \pi\delta(\omega - \beta) + \pi\delta(\omega + \beta)$ $e^{-\alpha\tau^2} \leftrightarrow \sqrt{\frac{\pi}{\alpha}} e^{-\omega^2/4\alpha}$
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Name: \_\_\_\_\_

ID #: \_\_\_\_\_

Directions:

1. **Print** your name and student number on the cover page.
2. Exam is closed book, closed notes, and no calculators.
3. Clearly designate all answers asked for (arrows, underline, box, etc.)

\_\_\_\_\_

ECE 600, Final Exam

Name: \_\_\_\_\_

Problems 1-10 are multiple choice problems worth 5 points each. For each problem, write the letter corresponding to the best answer next to the problem number. Space is provided to work out your solution for each of these problems. Please show your work! If your final grade is near a borderline, the quality of your written solutions could significantly impact your final grade.

1. Answer:

C

$$\begin{aligned}
 \Phi(\omega) &= E[e^{j\omega X}] \\
 &= \sum_{k=0}^n e^{j\omega k} \binom{n}{k} p^k (1-p)^{n-k} \\
 &= \sum_{k=0}^n \binom{n}{k} (pe^{j\omega})^k (1-p)^{n-k} \\
 &= \text{Binomial}(n, pm) \\
 &= (pe^{j\omega} + 1 - p)^n \\
 &= (1 + p(e^{j\omega} - 1))^n \\
 &\text{etc.}
 \end{aligned}$$

2. Answer:



11. Problems 11 is made up of 8 True/False questions, worth 2 points each. Fill in your answers T (true) or F (false) below, corresponding to the statements A-H in problem 11 on the exam.

A. \_\_\_\_\_

B. \_\_\_\_\_

C. \_\_\_\_\_

D. \_\_\_\_\_

E. \_\_\_\_\_

F. \_\_\_\_\_

G. \_\_\_\_\_

H. \_\_\_\_\_

*ECE 600, Final Exam*

Name: \_\_\_\_\_

Problems 12 and 13 are “work out” problems for which partial credit will be awarded for correctly reasoned work. It is important that you coherently present your thinking in the solution of these problems if you wish to receive partial credit (or full credit for that matter.) Please work problems 12 and 13 of the exam in the designated space below.

12. Problem 12 Solution:

*ECE 600, Final Exam*

Name: \_\_\_\_\_

13. Problem 13 Solution:

Thank You!

