

Session 29

Recall...

29.1

Example: Consider a R.P.

$$X(t) = \cos(\omega_0 t + \Theta)$$

where Θ is a R.V. uniformly distributed on $[0, 2\pi)$:

$$f_{\Theta}(\theta) = \begin{cases} \frac{1}{2\pi} & , 0 \leq \theta < 2\pi \\ 0 & , \text{elsewhere} \end{cases} , \omega_0 = \text{constant} \\ \text{(radian frequency, not outcome)}$$

Is $X(t)$ W.S.S? Let's check the two defining conditions

Recall...

$$(i) E[X(t)] = E[\cos(\omega_0 t + \Theta)]$$

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$$= E[\cos(\omega_0 t)\cos(\Theta) - \sin(\omega_0 t)\sin(\Theta)]$$

$$= \int_{-\infty}^{\infty} (\cos(\omega_0 t)\cos\theta - \sin(\omega_0 t)\sin(\theta)) f_{\Theta}(\theta) d\theta$$

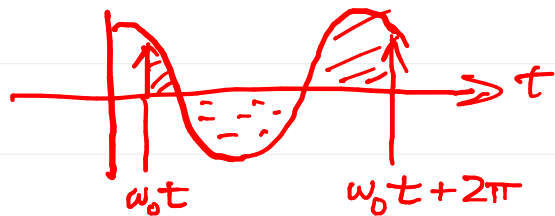
$$= \frac{\cos \omega_0 t}{2\pi} \int_0^{2\pi} \cos \theta d\theta - \frac{\sin \omega_0 t}{2\pi} \int_0^{2\pi} \sin \theta d\theta$$

$$= 0 \quad \therefore E[X(t)] = 0 = \text{constant}. \quad \checkmark$$

n.b. Even easier:

$$E[X(t)] = E[\cos(\omega_0 t + \Theta)] = \frac{1}{2\pi} \int_0^{2\pi} \cos(\omega_0 t + \theta) d\theta$$

Integral over one period is 0.



Recall...

$$\cos A \cdot \cos B = \frac{1}{2} [\cos(A+B) + \cos(A-B)]$$

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$$(ii) : E[X(t_1)X(t_2)]$$

$$= E[\cos(\omega_0 t_1 + \Theta) \cdot \cos(\omega_0 t_2 + \Theta)]$$

$$= E\left[\frac{1}{2} (\cos(\omega_0(t_1+t_2) + 2\Theta) + \cos(\omega_0(t_1-t_2)))\right]$$

$$= \frac{1}{2} E[\cos(\omega_0(t_1+t_2) + 2\Theta)] + \frac{1}{2} E[\cos(\omega_0(t_1-t_2))]$$

$$= \frac{1}{2} \cos(\omega_0(t_1-t_2)) \quad \checkmark$$

Recall...

$$\therefore (i) E[X(t)] = 0 = \text{constant},$$

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$$(ii) E[X(t_1)X(t_2)] = \frac{1}{2} \cos \omega_0 (t_1 - t_2).$$

$\therefore X(t)$ is a W.S.S. R.P.

Another Example:

$$\text{Let } \Psi(t) = \cos(\omega_0 t + \psi)$$

where $\omega_0 = \text{constant}$

and ψ is a uniform R.V. on $[0, \pi)$.

Is $\Psi(t)$ W.S.S.?

$$(i) E[\Psi(t)] = E[\cos(\omega_0 t + \psi)]$$

$$= \frac{1}{\pi} \int_0^{\pi} \cos(\omega_0 t + \psi) d\psi$$

$$= \frac{1}{\pi} \int_0^{\pi} [\cos \omega_0 t \cos \psi - \sin \omega_0 t \sin \psi] d\psi$$

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$$= \frac{\cos \omega_0 t}{\pi} \int_0^\pi \cos \psi d\psi - \frac{\sin \omega_0 t}{\pi} \int_0^\pi \sin \psi d\psi$$

$$= \frac{\sin \omega_0 t}{\pi} [-\cos \psi]_0^\pi = -\frac{2}{\pi} \sin \omega_0 t$$

\neq constant

$\therefore E[\psi(t)] \neq$ constant

$\Rightarrow \psi(t)$ is not W.S.S.

Defn: The mean of a R.P. $X(t)$
is

$$\mu_X(t) \triangleq E[X(t)]$$

Defn: The autocorrelation function
of a R.P. $X(t)$ is

$$R_{XX}(t_1, t_2) \triangleq E[X(t_1)X(t_2)]$$

n.b. The autocorrelation function
is just a measure of
the correlation in $X(t)$ at
time t_1 and time t_2 .

n.b. $R_{xx}(t_1, t_2) = E[x(t_1)x(t_2)]$

$$= E[x(t_2)x(t_1)] = R_{xx}(t_2, t_1)$$

for a real random process.

n.b. The autocorrelation function is symmetric in its time arguments.

For a complex random process $x(t)$

$$\begin{aligned} R_{xx}(t_1, t_2) &= E[x(t_1)x(t_2)^*] = (E[x(t_2)x(t_1)^*])^* \\ &= R_{xx}^*(t_2, t_1) \end{aligned}$$

Defn: The autocovariance
function of a ^{real} random process
 $X(t)$ is defined as

$$C_{XX}(t_1, t_2) \triangleq E[(X(t_1) - \mu_X(t_1))(X(t_2) - \mu_X(t_2))]$$

$$\left(= \dots = R_{XX}(t_1, t_2) - \mu_X(t_1)\mu_X(t_2) \right)$$

can easily show this

Defn: A random process $X(t)$ is called a Gaussian random process if the RVs $X(t_1), X(t_2), \dots, X(t_n)$ are jointly Gaussian for any $n \in \mathbb{N}$ and any set of sample times t_1, \dots, t_n .

Fact: The n -th order characteristic function of a Gaussian R.P. $X(t)$ is

$$\Phi_{X(t_1), \dots, X(t_n)}(w_1, \dots, w_n) = \exp \left\{ i \sum_{j=1}^n \mu_X(t_j) w_j \right\} \exp \left\{ -\frac{1}{2} \sum_{j=1}^n \sum_{k=1}^n C_{XX}(t_j, t_k) w_j w_k \right\}$$

This is a complete probabilistic description.

Fact: A Gaussian random process $X(t)$ is completely characterized by

(i) $\mu_X(t) = E[X(t)]$

and

(ii) $C_{XX}(t_1, t_2) = E[(X(t_1) - \mu_X(t_1))(X(t_2) - \mu_X(t_2))]$

Note: If a R.P $X(t)$ is W.S.S.,
then

$$\begin{aligned}R_{XX}(t_1, t_2) &= E[X(t_1)X(t_2)] \\ &= f(t_1 - t_2) \\ &= "R_X(t_1 - t_2)"\end{aligned}$$

This is also sometimes written
as

$$E[X(t+\tau)X(t)] = R_X(\tau)$$

for a W.S.S. random process.

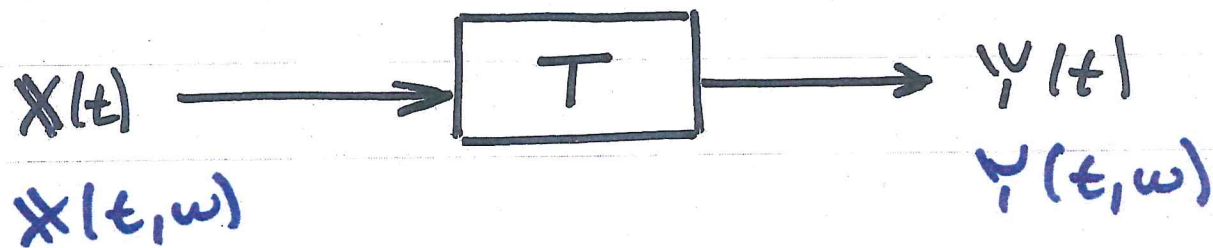
Defn: If $X(t)$ is a W.S.S. random process with autocorrelation function $R_x(\tau)$, then the Power Spectral Density of $X(t)$ is defined as

$$S_{XX}(\omega) \triangleq \int_{-\infty}^{\infty} R_x(\tau) e^{-i\omega\tau} d\tau$$

Here ω is radian frequency (not the outcome of experiment.)
 $S_{XX}(\omega)$ is a measure of the average distribution of signal power in frequency for the R.P. $X(t)$.

Systems with Stochastic Inputs

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Given a R.P. $X(t)$, if we assign to each sample function $X(t, \omega)$ a new sample function $Y(t, \omega)$, we have a new random process

$$Y(t) = T[X(t)]$$

whose sample functions are

$$Y(t, \omega) = T[X(t, \omega)].$$

n.b. We will assume that $T[\cdot]$ is deterministic. (not random)

Think of

$x(t)$ = input to a system

$y(t)$ = output of the system



$$y(t, \omega) = T[x(t, \omega)], \forall \omega \in \Omega.$$

- We are interested in finding a statistical description of the output $Y(t)$ in terms of the statistical description of the input $X(t)$ and the system description $T[\cdot]$
- For general T this is very difficult
- We will look at two special cases:
 1. Linear Time Invariant (L.T.I.) Systems
 2. Memoryless Systems.

Linear Systems

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A linear system $L[\cdot]$ is a transformation rule satisfying the following two properties:

$$1. L[x_1(t) + x_2(t)] = L[x_1(t)] + L[x_2(t)]$$

(Superposition)

$$2. L[A \cdot x(t)] = A \cdot L[x(t)]$$

(Homogeneity)

n.b. A can be an R.V. or a constant

Defn: A (linear) system is time-invariant if, given response $y(t)$ to input $x(t)$ it has response $y(t+c)$ for input $x(t+c)$, for all $c \in \mathbb{R}$.

A linear time invariant system is characterized by its impulse response $h(t)$:

$$\delta(t) \xrightarrow{\delta(t-t_0)} \boxed{h(t)} \longrightarrow h(t) \quad h(t-t_0)$$

$$x(t) \xrightarrow{\quad} \boxed{h(t)} \longrightarrow y(t) = x(t) * h(t).$$

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If we put a random process $X(t)$ into a L.T.I. system, we get a random process $Y(t)$ out of the system:

$$\begin{aligned} Y(t) &= X(t) * h(t) = \int_{-\infty}^{\infty} X(t-\alpha) h(\alpha) d\alpha \\ &= \int_{-\infty}^{\infty} X(\alpha) h(t-\alpha) d\alpha \end{aligned}$$

We interpret this on a sample function basis:

$$Y(t, \omega) = X(t, \omega) * h(t), \quad \forall \omega \in \Omega.$$

Important Facts:

1. If the input to a L.T.I. system is a Gaussian R.P., then the output is a Gaussian R.P.
2. If the input to a stable L.T.I. system is S.S.S., so is the output.

An L.T.I. system is stable iif
$$\int_{-\infty}^{\infty} |h(t)| dt < \infty.$$
(BIBO stable)

Fundamental Theorem:

For any linear system

$$E [L [x(t)]] = L [E [x(t)]]$$

(This basically reduces to an exchange of orders of integration).

Applying this to a L.T.I. system, we get

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$$\begin{aligned} E[Y(t)] &= E\left[\int_{-\infty}^{\infty} X(t-\alpha)h(\alpha) d\alpha\right] \\ &= \int_{-\infty}^{\infty} E[X(t-\alpha)]h(\alpha) d\alpha \\ &= \int_{-\infty}^{\infty} \mu_X(t-\alpha)h(\alpha) d\alpha \\ &= \mu_X(t) * h(t) \end{aligned}$$

$$\therefore \mu_Y(t) = E[Y(t)] = \mu_X(t) * h(t).$$

$$R_{yy}(t_1, t_2) = E [Y(t_1) Y(t_2)]$$

$$= E \left[\underbrace{\int_{-\infty}^{\infty} X(t_1 - \alpha) h(\alpha) d\alpha}_{Y(t_1)} \cdot \underbrace{\int_{-\infty}^{\infty} X(t_2 - \beta) h(\beta) d\beta}_{Y(t_2)} \right]$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} E [X(t_1 - \alpha) X(t_2 - \beta)] h(\alpha) h(\beta) d\alpha d\beta$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} R_{xx}(t_1 - \alpha, t_2 - \beta) h(\alpha) h(\beta) d\alpha d\beta$$

If $R_{xx}(t_1, t_2) = R_x(t_1 - t_2)$, then we would have

$$R_{yy}(t_1, t_2) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} R_x(\underbrace{t_1 - t_2 - \alpha + \beta}_{\downarrow}) h(\alpha) h(\beta) d\alpha d\beta$$

$$= R_y(t_1 - t_2) \quad (\text{a function of } t_1 - t_2)$$

So if $X(t)$ is a W.S.S.
random process, then

$$\begin{aligned} \text{(i) } \mu_Y(t) &= \mu_X * h(t) = \int_{-\infty}^{\infty} \mu_X h(t-\alpha) d\alpha \\ &= \mu_X \int_{-\infty}^{\infty} h(t-\alpha) d\alpha = \mu_Y \end{aligned}$$

$$\text{(ii) } R_{YY}(t_1, t_2) = R_Y(t_1 - t_2)$$

Theorem: If the input to a stable L.T.I.
system is a W.S.S. random process,
then the output is a W.S.S.
random process.

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Recall the definition of
 the Power Spectral Density of a W.S.S

R.P. $X(t)$ is

$$S_{XX}(\omega) \triangleq \int_{-\infty}^{\infty} R_{XX}(\tau) e^{-i\omega\tau} d\tau$$

$\omega = \text{radian}$
 500°

(where we assume $X(t)$ is W.S.S.)

where

$$R_{XX}(\tau) = E \left[\underbrace{X(t+\tau)}_{t_1} \underbrace{X^*(t)}_{t_2} \right]$$

for a W.S.S. R.P. $X(t)$

Note: 1. Because $R_{xx}(-\tau) = R_{xx}^*(\tau)$

$S_{xx}(\omega)$ is a real function

$$2. R_{xx}(\tau) = \frac{1}{2\pi} \int_{-\infty}^{\infty} S_{xx}(\omega) e^{i\omega\tau} d\omega$$

3. In order to consider $S_{xx}(\omega)$, we must assume $x(t)$ is W.S.S.

4. Because $R_{xx}(\tau)$ is a non-negative definite function, it follows that $S_{xx}(\omega) \geq 0$.

Key Result: If $X(t)$ is a W.S.S.

Random Process and it is the input to a stable L.T.I. system with impulse response $h(t)$, then the Power Spectral Density of the output $Y(t)$ is

$$S_{YY}(\omega) = S_{XX}(\omega) |H(\omega)|^2$$

where

$$H(\omega) = \int_{-\infty}^{\infty} h(t) e^{-i\omega t} dt.$$

Defn : A random process $W(t)$ is called a white noise process if

$$C_{ww}(t_1, t_2) = 0, \forall t_1 \neq t_2.$$

Fact : All (non-trivial) W.S.S. white noise processes have

$$R_{ww}(t_1, t_2) = r_0 \cdot \delta(t_1 - t_2)$$

where $r_0 = \text{constant} > 0$.

29.29

EXAMPLE: Let $X(t)$ be a W.S.S

R.P. with PSD $S_{XX}(\omega)$ and let

$Y(t)$ be the "smoothed" random process given by

$$Y(t) = \frac{1}{2T} \int_{t-T}^{t+T} X(\alpha) d\alpha$$

This can be represented by a L.T.I. System with impulse response

$$h(t) = \frac{1}{2T} \cdot \mathbb{1}_{[-T, T]}(t)$$



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$$h(t) = \frac{1}{2T} \cdot \mathbf{1}_{[-T, T]}(t)$$

What is the PSD $S_{YY}(\omega)$ of $Y(t)$?

Solution:

$$S_{YY}(\omega) = |H(\omega)|^2 S_{XX}(\omega)$$

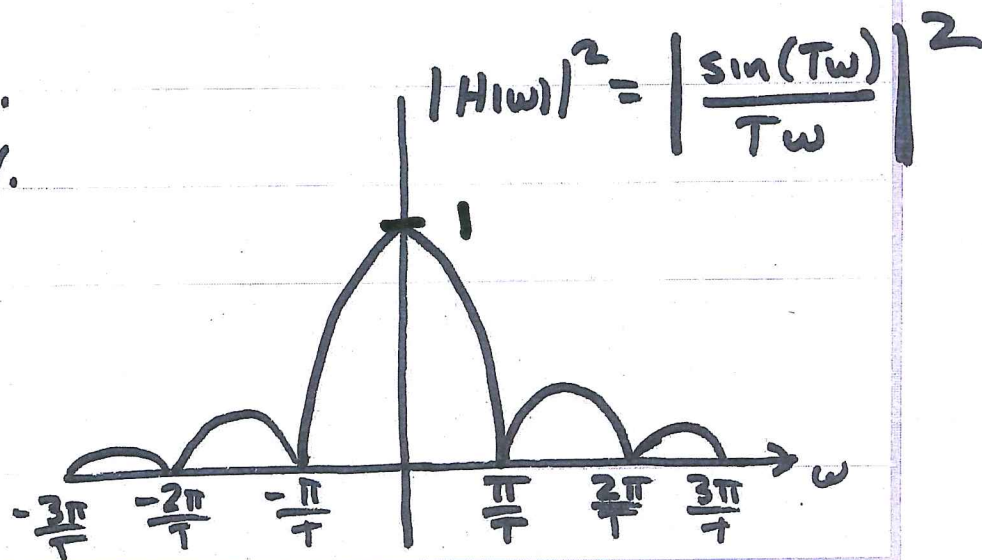
$$H(\omega) = \int_{-\infty}^{\infty} h(t) e^{-i\omega t} dt = \int_{-\infty}^{\infty} \frac{1}{2T} \mathbf{1}_{[-T, T]}(t) e^{-i\omega t} dt$$

$$= \frac{1}{2T} \int_{-T}^T e^{-i\omega t} dt = \frac{1}{2T} \left. \frac{e^{-i\omega t}}{-i\omega} \right|_{-T}^T$$

$$= \frac{e^{-i\omega T} - e^{+i\omega T}}{-i2T\omega} = \frac{1}{T\omega} \left[\frac{e^{i\omega T} - e^{-i\omega T}}{i2} \right] = \frac{\sin(T\omega)}{T\omega}$$

$$S_{YY}(\omega) = S_{XX}(\omega) \left| \frac{\sin(T\omega)}{T\omega} \right|^2$$

Note that $h(t)$ acts as a crude low-pass filter that attenuates high-freq. power in the signal.



Given a W.S.S RP $X(t)$ that
is the input to an L.T.I system:



If I want to find the autocorrelation function of the output I have two options:

$$(1) R_Y(z) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} R_X(z^{-\alpha+\beta}) h(\alpha) h(\beta) d\alpha d\beta$$

Most of the time this is a complicated calculation.

(2) Use $S_{yy}(\omega) = |H(\omega)|^2 \cdot S_{xx}(\omega)$

(a) Compute

$$S_{xx}(\omega) = \int_{-\infty}^{\infty} R_x(\tau) e^{-i\omega\tau} d\tau$$

(b) Compute

$$H(\omega) = \int_{-\infty}^{\infty} h(\tau) e^{-i\omega\tau} d\tau$$

(c) Compute

$$S_{yy}(\omega) = S_{xx}(\omega) \cdot |H(\omega)|^2$$

(d) Compute

$$R_y(\tau) = \frac{1}{2\pi} \int_{-\infty}^{\infty} S_{yy}(\omega) e^{+i\omega\tau} d\omega$$

Often the inverse Fourier transform can be looked up in a table:

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TABLE 10-1*

$$R(\tau) = \frac{1}{2\pi} \int_{-\infty}^{\infty} S(\omega) e^{j\omega\tau} d\omega \leftrightarrow S(\omega) = \int_{-\infty}^{\infty} R(\tau) e^{-j\omega\tau} d\tau$$

$$\delta(\tau) \leftrightarrow 1$$

$$1 \leftrightarrow 2\pi\delta(\omega)$$

$$e^{j\beta\tau} \leftrightarrow 2\pi\delta(\omega - \beta)$$

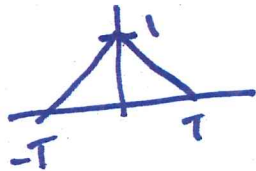
$$\cos \beta\tau \leftrightarrow \pi\delta(\omega - \beta) + \pi\delta(\omega + \beta)$$

$$e^{-\alpha|\tau|} \leftrightarrow \frac{2\alpha}{\alpha^2 + \omega^2}$$

$$e^{-\alpha\tau^2} \leftrightarrow \sqrt{\frac{\pi}{\alpha}} e^{-\omega^2/4\alpha}$$

$$e^{-\alpha|\tau|} \cos \beta\tau \leftrightarrow \frac{\alpha}{\alpha^2 + (\omega - \beta)^2} + \frac{\alpha}{\alpha^2 + (\omega + \beta)^2}$$

$$2e^{-\alpha\tau^2} \cos \beta\tau \leftrightarrow \sqrt{\frac{\pi}{\alpha}} [e^{-(\omega - \beta)^2/4\alpha} + e^{-(\omega + \beta)^2/4\alpha}]$$

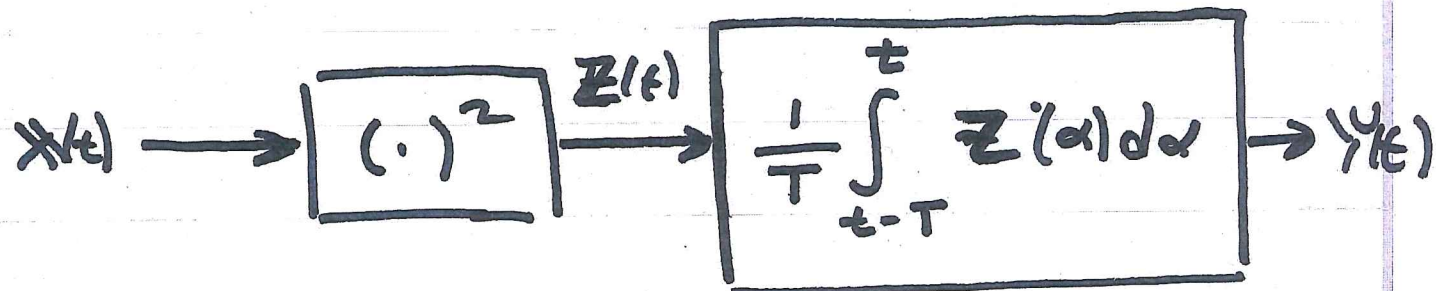


$$\begin{cases} 1 - \frac{|\tau|}{T} & |\tau| < T \\ 0 & |\tau| > T \end{cases} \leftrightarrow \frac{4 \sin^2(\omega T/2)}{T\omega^2}$$

$$\frac{\sin \sigma\tau}{\pi\tau} \leftrightarrow \begin{cases} 1 & |\omega| < \sigma \\ 0 & |\omega| > \sigma \end{cases}$$

* This table will be included in the final exam.

Aside: A common way to estimate the instantaneous power in a signal is to cascade a memoryless system and a L.T.I. System



L.T.I. System
with

$$h(t) = \frac{1}{T} \mathbf{1}_{[0, T]}(t)$$

Memoryless Systems

29.36

Defn : A system is called memoryless if its output at time t is

$$Y(t) = T[X(t)] = g(X(t)),$$

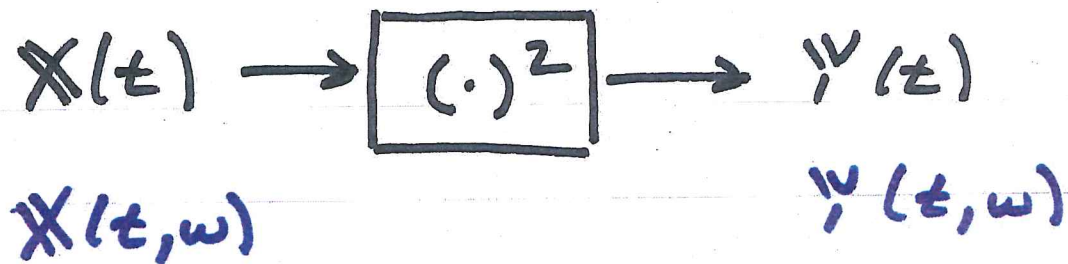
where $g: \mathbb{R} \rightarrow \mathbb{R}$, is only a function only of the current value of $X(t)$.

- $g(\cdot)$ is not a function of past or future values of its input
- $Y(t) = g(X(t))$ depends only on the instantaneous value of $X(t)$ at time t .

Ex. 1. $g(x) = x^2$

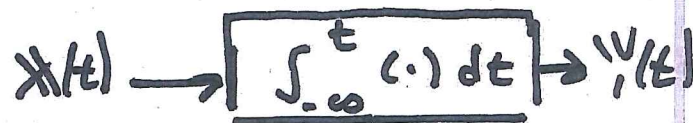
$$Y(t) = g(X(t)) = X^2(t)$$

is a memoryless system.



Ex. 2 Integrators are not memoryless. They have memory of the past.

$$Y(t) = \int_{-\infty}^t X(\alpha) d\alpha$$



$$Y(t, \omega) = \int_{-\infty}^t X(\alpha, \omega) d\alpha$$

For a memoryless system
with

$$Y(t) = g(X(t)),$$

the first order density $f_{Y(t)}(y)$ of
 $Y(t)$ can be expressed in terms
of the first order density $f_{X(t)}(x)$
of $X(t)$ and $g(\cdot)$

$X(t)$ is just a R.V. and
 $Y(t) = g(X(t))$ is just a
transformed RV (function of $X(t)$.)

Note also that

$$E[Y(t)] = E[g(X(t))] = \int_{-\infty}^{\infty} g(x) f_{X(t)}(x) dx$$

and

$$R_{YY}(t_1, t_2) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x_1) g(x_2) \cdot f_{X(t_1), X(t_2)}(x_1, x_2) dx_1 dx_2$$

Also, you can get the n -th order pdf of $Y(t)$ using the mapping

$$Y(t_1) = g(X(t_1)), Y(t_2) = g(X(t_2)), \\ \dots, Y(t_n) = g(X(t_n)).$$

Theorem: Let $x(t)$ be a S.S.S.
R.P. that is the input to
a memoryless system. Then
the output is also a S.S.S.
R.P.

Proof: Papoulis P. 394 (P. 304 in 3rd Ed.)

$$f_{y_1(t+c), \dots, y_n(t+c)}^{(y_1, \dots, y_n)} = f_{x_1(y), \dots, x_n(y)}^{(x_1, \dots, x_n)} \left| \frac{\partial(x_1, \dots, x_n)}{\partial(y_1, \dots, y_n)} \right|$$

$$= f_{x_1(y), \dots, x_n(y)}^{(x_1, \dots, x_n)} \left| \frac{\partial(x_1, \dots, x_n)}{\partial(y_1, \dots, y_n)} \right| = f_{y_1(t), \dots, y_n(t)}^{(y_1, \dots, y_n)}$$

$\forall c \in \mathbb{R}, \forall n \in \mathbb{N}$ and all t_1, \dots, t_n .

not a fn.
of c .

n.b. Regarding the invariance of the Jacobian with
a change in time origin;
Assuming $g(\cdot)$ is invertible and $h(\cdot) = g^{-1}(\cdot)$

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$$\begin{aligned} x_1 &= h(y_1) \\ x_2 &= h(y_2) \\ &\vdots \\ x_n &= h(y_n) \end{aligned} \Rightarrow \frac{\partial(x_1, \dots, x_n)}{\partial(y_1, \dots, y_n)} = \begin{pmatrix} \frac{\partial h(y_1)}{\partial y_1} & 0 & 0 & \dots & 0 \\ 0 & \frac{\partial h(y_2)}{\partial y_2} & 0 & \dots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \dots & \frac{\partial h(y_n)}{\partial y_n} \end{pmatrix}$$
$$= \frac{\partial h(y_1)}{\partial y_1} \cdot \frac{\partial h(y_2)}{\partial y_2} \dots \frac{\partial h(y_n)}{\partial y_n}$$

Since $g(x)$ does not vary with time, neither
does $h(y) = g^{-1}(y)$, so the Jacobian does not
depend on time.

Example: Hard Limiter

Consider a memoryless system with

$$g(x) = \begin{cases} +1, & x \geq 0, \\ -1, & x < 0 \end{cases}$$

Consider $Y(t) = g(X(t)) = \text{sgn}(X(t))$

Find $E[Y(t)]$ and $R_{YY}(t_1, t_2)$ given the "statistics" of $X(t)$.

$$E[Y(t)] = (+1) \cdot P(\{Y(t) = 1\})$$

$$+ (-1) \cdot P(\{Y(t) = -1\})$$

$$= (+1) \cdot P(\{X(t) \geq 0\}) + (-1) \cdot P(\{X(t) < 0\})$$

$$= 1 - F_{X(t)}(0) - F_{X(t)}(0)$$

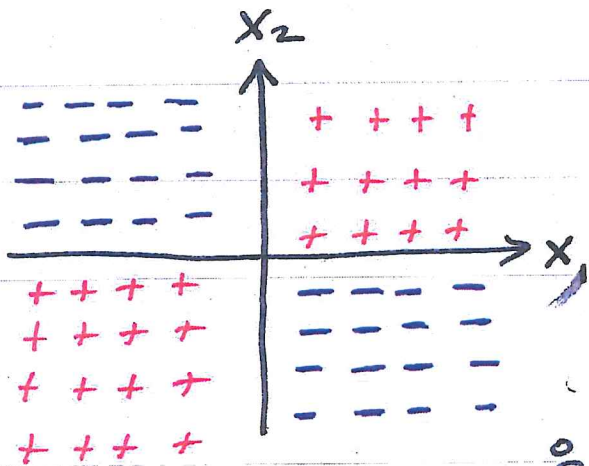
$$= 1 - 2F_{X(t)}(0).$$

$$R_{\psi\psi}(t_1, t_2) = E[\psi(t_1) \cdot \psi(t_2)]$$

note: $\psi(t_1) \cdot \psi(t_2) = \begin{cases} +1, & \psi(t_1) \psi(t_2) > 0 \\ -1, & \psi(t_1) \psi(t_2) < 0 \end{cases}$

So

$$E[\psi(t_1) \psi(t_2)] = 1 \cdot P(\{\psi(t_1) \psi(t_2) \geq 0\}) - 1 \cdot P(\{\psi(t_1) \psi(t_2) < 0\})$$



Where

$$P(\{\psi(t_1) \psi(t_2) \geq 0\}) = \int_0^{\infty} \int_0^{\infty} f(x_1, x_2) dx_1 dx_2 + \int_{-\infty}^0 \int_{-\infty}^0 f(x_1, x_2) dx_1 dx_2$$

und $P(\{\psi(t_1) \psi(t_2) < 0\}) = \int_{-\infty}^0 \int_0^{\infty} f(x_1, x_2) dx_1 dx_2 + \int_0^{\infty} \int_{-\infty}^0 f(x_1, x_2) dx_1 dx_2$