

Session 23

Iterated Expectation

23.1

Another common situation:

$$E[g(X, Y)] = \iint_{\mathbb{R}^2} g(x, y) f_{X,Y}(x, y) dx dy$$

$$= \iint_{\mathbb{R}^2} g(x, y) f_{Y|X}(y|x) f_X(x) dx dy$$

$$= \int_R f_X(x) \underbrace{\left[\int_R g(x, y) f_{Y|X}(y|x) dy \right]}_{\varphi(x)} dx .$$

$$= \int_R f_X(x) E[g(x, Y) | X=x] dx = \dots$$

$$(= E[\varphi(X)])$$

23.2

$$\dots = \int_{-\infty}^{\infty} f_{\mathbb{X}}(x) \underbrace{E[g(\mathbb{X}, \gamma) | \mathbb{X}=x]}_{\varphi(x)} dx$$

$$= \int_{-\infty}^{\infty} f_{\mathbb{X}}(x) \varphi(x) dx = E_{\mathbb{X}}[\varphi(\mathbb{X})] \leftarrow$$

where $\varphi(x) = E[g(\mathbb{X}, \gamma) | \{\mathbb{X}=x\}]$

$$\begin{aligned} \therefore E[g(\mathbb{X}, \gamma)] &= E_{\mathbb{X}}[\varphi(\mathbb{X})] \\ &= E_{\mathbb{X}}[E_{\gamma} [g(\mathbb{X}, \gamma) | \mathbb{X}]] \\ &= E[E[g(\mathbb{X}, \gamma) | \mathbb{X}]] \end{aligned}$$

note the notation ↪

Summarizing, we have

23.3

$$\underset{xy}{E}[g(x,y)] = \underset{x}{E}\left[\underset{y}{E}[g(x,y)|x]\right]$$

n.b. The terminology "iterated" comes from
"integrated integration"

$$E[g(x,y)] = \int_{-\infty}^{\infty} f_x(x) \int_{-\infty}^{\infty} g(x,y) f_y(y | \{x=x\}) dy dx$$



$$E_y [g(x,y) | \{x=x\}]$$

One very important application of iterated expectation is Minimum Mean-Square estimation

So we have

$$E[g(X, Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f_{X,Y}(x, y) dx dy$$

$$= \dots = \int_{-\infty}^{\infty} f_X(x) \int_{-\infty}^{\infty} g(x, y) f_{Y|X}(y | \{X=x\}) dy dx$$

$E[g(X, Y) | \{X=x\}] = \varphi(x)$

$$= E_x [E_y [g(X, Y) | X]] = E_x [\varphi(X)].$$

Minimum Mean-Square Estimation

23.5

- Let X and Y be two j-dist RVs with j-pdf $f_{XY}(x, y)$.
- Suppose we want to estimate the value of Y given that we have observed $\{\sum X = x\}$.

Q: What is the best estimate of the value of Y given that we know $X=x$?

What do we mean by best?

One commonly used error criterion is square error.
⇒ Design the estimator that minimizes mean-square-error.

23.6

We want to find the function $c(x)$ to estimate the value of y given that we observe $\hat{x}=x$ such that

$$\mathcal{E} = E[(y - c(\hat{x}))^2]$$

is minimized.

Claim : The mean-square error \mathcal{E} is minimized when

$$c(x) = E[y | \{\hat{x}=x\}].$$

$$\text{Proof: } \mathcal{E} = E \left[(\underline{y} - c(\underline{x}))^2 \right]$$

$$\begin{aligned}
 &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (\underline{y} - c(x))^2 f_{\underline{x}, \underline{y}}(x, y) dx dy \\
 &= \int_{-\infty}^{\infty} f_x(x) \left[\int_{-\infty}^{\infty} (\underline{y} - c(x))^2 \cdot f_y(y|x) dy \right] dx
 \end{aligned}$$

$\underbrace{\quad}_{\varphi(x)}$

\mathcal{E} is minimized if we pick $c(x)$ such that the inner integral is minimized for each value of x .

23.8

We minimize the inner integral
as follows (for any particular x):

$$\begin{aligned} \frac{\partial}{\partial c(x)} & \left\{ \int_{-\infty}^{\infty} [\gamma - c(x)]^2 f_{\gamma|X}(\gamma|x) d\gamma \right\} \\ &= -2 \int_{-\infty}^{\infty} [\gamma - c(x)] f_{\gamma|X}(\gamma|x) d\gamma = 0 \end{aligned}$$

$$\Rightarrow \int_{-\infty}^{\infty} (\gamma - c(x)) f_{\gamma|X}(\gamma|x) d\gamma = 0$$

$$\Rightarrow \underbrace{\int_{-\infty}^{\infty} \gamma f_{\gamma|X}(\gamma|x) d\gamma}_{E[\gamma|X=x]} - c(x) \underbrace{\int_{-\infty}^{\infty} f_{\gamma|X}(\gamma|x) d\gamma}_1 =$$

$$E[\gamma|X=x] - c(x) = 0$$

$$\Rightarrow c(x) = E[Y | X=x]$$

23.9

We will use the following notation
for \hat{Y}^k MMSE estimator

$$\hat{Y}_{\text{MMSE}}(x) = E[Y | X=x]$$

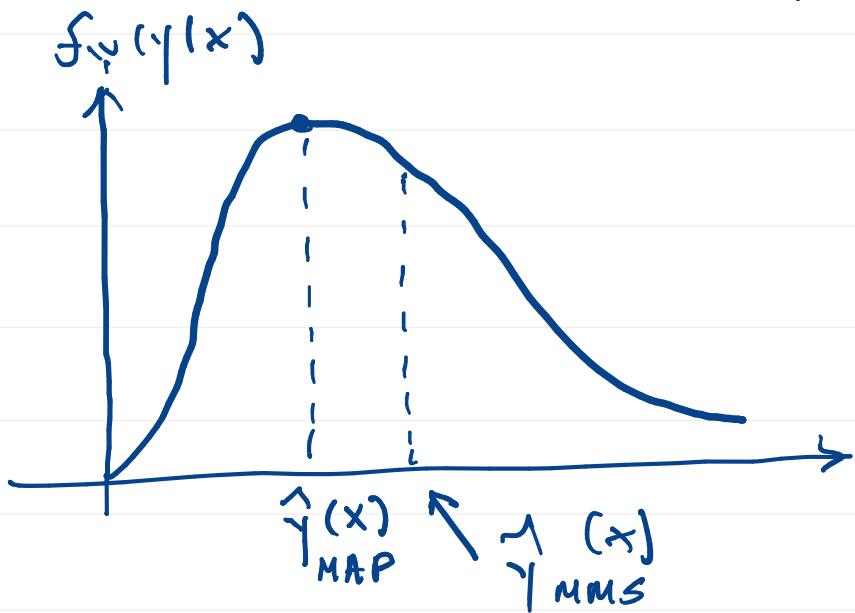
Similarly, by symmetry

$$\hat{X}_{\text{MMSE}}(y) = E[X | Y=y] \quad (\text{Exercise})$$

Consider another estimator:

23.10

$$\hat{y}_{MAP}(x) = \arg \max_y \{ f_{\hat{y}}(y|x) \}$$



MAP \triangleq Maximum
Aposteriori
Probability/

The MAP estimators of interest are

$$\hat{y}_{MAP}(x) = \arg \max_y \{ f_{\hat{y}}(y|x=x) \}$$

$$\hat{x}_{MAP}(y) = \arg \max_x \{ f_x(x|y=y) \}.$$

Papoulis in the reading discussed

the Linear Minimum Mean Square Error
(LMMSE) estimator

$$c(x) = \underline{ax + b} \quad \} \text{ assumed form.}$$

Don't worry about this. You
are not responsible for it

Random Vectors

23.12

- We've considered two RVs on $(\mathcal{S}, \mathcal{F}, P)$.
 - We can extend this to n RVs on $(\mathcal{S}, \mathcal{F}, P)$:
- $X_1(\omega), X_2(\omega), \dots, X_n(\omega).$
- We can arrange these RVs as elements of a vector.

(A random vector.)

Defn: Let X_1, \dots, X_n be n jointly distributed RVs on $(\mathcal{S}, \mathcal{F}, P)$. Then the vector of RVs

$$\underline{X} = (X_1, \dots, X_n) \xleftarrow{\text{Row Vector}}$$

is a random vector (RVec)
defined on $(\mathcal{S}, \mathcal{F}, P)$.

Alternatively, we can think of a RVec as a mapping from \mathcal{S} to \mathbb{R}^n

$$\underline{X} : \mathcal{S} \rightarrow \mathbb{R}^n$$

23.14

$$\underline{\text{CDF}}: F_{\underline{\underline{X}}}(\underline{x}) = F_{\underline{\underline{X}}}(\underline{x}_1, \dots, \underline{x}_n)$$

$$= F_{\underline{\underline{X}}_1, \dots, \underline{\underline{X}}_n}(\underline{x}_1, \dots, \underline{x}_n)$$

$$= P(\{\underline{X}_1 \leq \underline{x}_1\} \cap \{\underline{X}_2 \leq \underline{x}_2\} \cap \dots \cap \{\underline{X}_n \leq \underline{x}_n\})$$

$$\underline{\text{PDF}}: f_{\underline{\underline{X}}}(\underline{x}) = f_{\underline{\underline{X}}}(\underline{x}_1, \dots, \underline{x}_n)$$

$$= f_{\underline{\underline{X}}_1, \dots, \underline{\underline{X}}_n}(\underline{x}_1, \dots, \underline{x}_n)$$

$$= \frac{\partial^n}{\partial x_1 \partial x_2 \cdots \partial x_n} F_{\underline{\underline{X}}}(\underline{x}_1, \dots, \underline{x}_n)$$

23.15

Let $D \subset \mathbb{R}^n$ ($D \in \mathcal{B}(\mathbb{R}^n)$)

$$\text{Then } P(\{\underline{x} \in D\}) = \int_D f_{\underline{x}}(\underline{x}) \, d\underline{x}$$

$$= \int_{\mathbb{R}^n} f_{\underline{x}}(\underline{x}) \cdot 1_D(\underline{x}) \, d\underline{x}$$

$$= \underbrace{\int_{-\infty}^{\infty} \cdots \int_{-\infty}^{\infty}}_{n\text{-fold integration}} f_{\underline{x}}(x_1, \dots, x_n) \cdot 1_D((x_1, \dots, x_n)) \, dx_1 \cdots dx_n$$

23.16

All the general properties of
 cdfs and pdfs generalize from
 2-dimensions to n-dimensions:

e.g. Given j-dist RVs X_1, X_2, X_3, X_4

$$F_{X_1, X_3}(x_1, x_3) = F_{X_1, +\infty, X_3, +\infty}(x_1, +\infty, x_3, +\infty).$$

or

$$f_{X_1, X_3}(x_1, x_3) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X_1, X_2, X_3, X_4}(x_1, x_2, x_3, x_4) dx_2 dx_4$$

Transformations on RVecs

27.17

Given $\underline{x} = (x_1, \dots, x_n) \sim R\text{Vec}$

we can define a new RVec

$$\underline{y} = (y_1, \dots, y_k)$$

where $y_j = g_j(\underline{x})$, $j = 1, \dots, k$

where $k \leq n$

How do we find $F_{\underline{y}}(y)$ or $f_{\underline{y}}(y)$?

23.18

Given $\underline{X} = (X_1, \dots, X_n) \sim R\text{Vec}$

we can define the new $R\text{Vec}$

$$\underline{Y} = (Y_1, \dots, Y_k),$$

where

$$Y_j = g_j(\underline{X}), \quad j = 1, \dots, k.$$

Here $k \leq n$.

How do we find $F_{\underline{Y}}(y)$ or $f_{\underline{Y}}(y)$?

23.1

23.19

To find $F_{\underline{Y}}(y_1, \dots, y_k)$, define

$$D(y_1, \dots, y_k) \triangleq \left\{ (x_1, \dots, x_n) \in \mathbb{R}^n : g_1(x_1, \dots, x_n) \leq y_1, \dots, g_k(x_1, \dots, x_n) \leq y_k \right\} \subset \mathbb{R}^n (\in \mathcal{B}(\mathbb{R}^n))$$

Then

$$F_{\underline{Y}}(y_1, \dots, y_k) = \int_{D(y_1, \dots, y_k)} \cdots \int f_{\underline{X}}(x_1, \dots, x_n) dx_1 \cdots dx_n$$

and

$$f_{\underline{Y}}(y_1, \dots, y_k) = \frac{\partial^k F_{\underline{Y}}(y_1, \dots, y_k)}{\partial y_1 \partial y_2 \cdots \partial y_k}.$$

Direct Approach ($k=n$ and one-to-one mapping)

23.20

Theorem: Given $R\text{Vec } \underline{\mathbb{X}} = (\mathbb{X}_1, \dots, \mathbb{X}_n)$,
define a new $R\text{Vec } \underline{\mathbb{Y}} = (\mathbb{Y}_1, \dots, \mathbb{Y}_n)$
by the one-to-one mapping

$$G : \mathbb{R}^n \rightarrow \mathbb{R}^n$$

described by

$$\mathbb{Y}_1 = g_1(\underline{\mathbb{X}}), \dots, \mathbb{Y}_n = g_n(\underline{\mathbb{X}}).$$

Let $\mathbb{X}_1 = h_1(\underline{\mathbb{Y}}) = x_1(\underline{\mathbb{Y}}), \dots,$

$$\mathbb{X}_n = h_n(\underline{\mathbb{Y}}) = x_n(\underline{\mathbb{Y}}).$$

23.21

Then the j-pdf of \underline{Y} is

$$f_{\underline{y}}(y_1, \dots, y_n) = f_{\underline{x}}(x_1(y), \dots, x_n(y)) \left| \frac{\partial(x_1, \dots, x_n)}{\partial(y_1, \dots, y_n)} \right|$$

where the Jacobian is given by

$$\left| \frac{\partial(x_1, \dots, x_n)}{\partial(y_1, \dots, y_n)} \right| = \begin{vmatrix} \frac{\partial x_1}{\partial y_1} & \frac{\partial x_1}{\partial y_2} & \dots & \frac{\partial x_1}{\partial y_n} \\ \frac{\partial x_2}{\partial y_1} & \frac{\partial x_2}{\partial y_2} & \dots & \frac{\partial x_2}{\partial y_n} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial x_n}{\partial y_1} & \frac{\partial x_n}{\partial y_2} & \dots & \frac{\partial x_n}{\partial y_n} \end{vmatrix}$$

n.b. If you have y_1, \dots, y_k , where $k < n$, you can introduce aux. variables.