

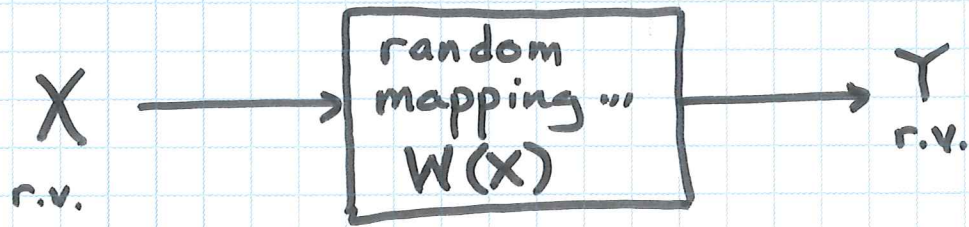
ECE 645

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Joint Distributions of R.V.s



Observation $Y=y$
Goal: Learn something about X .

\Rightarrow

$$F_{X,Y}(x,y) = P(X \leq x, Y \leq y)$$

Correlation, Covariance, joint pdf

$$F_{X|Y=y}(x|y) = P(X \leq x | Y=y)$$

$$= \int_{-\infty}^x f_{X|Y}(\alpha|y) d\alpha$$

$$f_{X,Y}(\alpha,y) / f_Y(y)$$

... $F_{X|Y}, f_{X|Y}$...

$P(X \in A | Y=y)$

...
mean, var cond. on $Y=y$

Important Props of Conditional Expectation

$$E\{X|Y=y\} \triangleq Q(y)$$

$$\triangleq \int_{-\infty}^{\infty} x f_{X|Y}(x|y) dx$$

X, Y jointly distributed ~~#~~ & $g(\cdot)$
a real-valued function s.t.

\Rightarrow can create a new
rv. by $Q \triangleq Q(Y)$

$$E\{|g(x)|\} < \infty$$

- Then :
- (a) $X \perp\!\!\!\perp Y \Rightarrow E\{g(x)|Y\} = E\{g(x)\}$
 - (b) X is a function of Y i.e. $X = h(Y)$
 $E\{g(x)|Y\} = E\{g(h(Y))|Y\} = g(h(Y))$
 - (c) $E\{g(x)\} = E\{E\{g(x)|Y\}\}$
 - (d) $E\{g(x)f(Y)|Y\} = E\{g(x)|Y\}f(Y)$

Proof © "Law of Total Expectation"

$$E\{g(x) | Y=y\} = \int_{-\infty}^{\infty} g(x) f_{X|Y}(x|y) dx \rightarrow \text{this is a funct. of } y. \text{ say } h(y).$$

If we subst. the rv. Y into $h(y)$ get a new rv $h(Y)$.

$$E\{h(Y)\} = \int_{-\infty}^{\infty} h(y) f_Y(y) dy = \int_{-\infty}^{\infty} \left[\int_{-\infty}^{\infty} g(x) f_{X|Y}(x|y) dx \right] f_Y(y) dy$$

Then use $f_{X,Y}(x,y) = f_{X|Y}(x|y) f_Y(y)$

$$\begin{aligned} \Rightarrow E\{h(Y)\} &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x) f_{X,Y}(x,y) dx dy = \int_{-\infty}^{\infty} g(x) \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy dx \\ &= \int_{-\infty}^{\infty} g(x) f_X(x) dx = E\{g(x)\}. \end{aligned}$$

Best Estimator Property

X, Y jointly dist. rvs and $g(\cdot)$ a measurable function (i.e., so $g(X)$ is also a r.v.)

Assume all needed expectations exist. Consider problem of estimating $g(X)$ via functions of Y

$$\hat{\phi} = \operatorname{argmin}_{\phi \text{ Borel}} E\left\{\left(g(X) - \phi(Y)\right)^2\right\}$$

Then the MSE is minimized over all Borel functions ϕ by choosing

$$\hat{\phi}(Y) = E\{g(X) | Y\}$$

Proof of B.E.P.

$$E\{(g(X) - \phi(Y))^2\} = E\{E\{(g(X) - \phi(Y))^2 | Y\}\}$$

$$= \int (g(x) - \phi(y))^2 f_{X|Y}(x|y) dx$$

$$= \int [\quad] f_X(y) dy$$

The double integral is minimized if the interior integral is minimized pointwise for each y , i.e. if we pick $\phi(y)$ st.

$E\{(g(X) - \phi(y))^2 | Y=y\}$ is minimized.

Back to here: $\hat{\phi}(y) = E\{g(X) | Y=y\}$.

because $f_{X|Y}(x|y)$ is a valid pdf for X .

Lemma Z is a
r.v. with finite mean

$$\hat{b} = \operatorname{argmin}_b E\{(Z-b)^2\}$$

$$\hat{b} = E\{Z\}.$$

Table G.1 A Short Table of Q-Function Values

x	$Q(x)$	x	$Q(x)$	x	$Q(x)$
0	0.5	1.5	0.066807	3.0	0.0013499
0.1	0.46017	1.6	0.054799	3.1	0.00096760
0.2	0.42074	1.7	0.044565	3.2	0.00068714
0.3	0.38209	1.8	0.035930	3.3	0.00048342
0.4	0.34458	1.9	0.028717	3.4	0.00033693
0.5	0.30854	2.0	0.022750	3.5	0.00023263
0.6	0.27425	2.1	0.017864	3.6	0.00015911
0.7	0.24196	2.2	0.013903	3.7	0.00010780
0.8	0.21186	2.3	0.010724	3.8	7.2348×10^{-5}
0.9	0.18406	2.4	0.0081975	3.9	4.8096×10^{-5}
1.0	0.15866	2.5	0.0062097	4.0	3.1671×10^{-5}
1.1	0.13567	2.6	0.0046612	4.1	2.0658×10^{-5}
1.2	0.11507	2.7	0.0034670	4.2	1.3346×10^{-5}
1.3	0.096800	2.8	0.0025551	4.3	8.5399×10^{-6}
1.4	0.080757	2.9	0.0018658	4.4	5.4125×10^{-6}

2.5.5 Conditional Expectation

+ Conditional expectation of X given $Y = y$ is ...

$$E[X|Y = y] = \begin{cases} \sum_x xp_{X|Y}(x|y) & \text{discrete} \\ \int x f_{X|Y}(x|y) dx & \text{continuous} \end{cases}$$

+ Law of Total Expectation. For rvs X, Y ...

$$E[X] = E[E[X|Y]].$$

2.6 Joint Gaussian Distribution

+ Two r.v.s X and Y are said to be bivariate normal (or Gaussian) if their joint pdf is

$$f_{X,Y}(x, y) = \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} \exp\left\{-\frac{1}{2(1-\rho^2)}F(x, y)\right\}$$

$$\left(\frac{x-\mu_x}{\sigma_x}\right)^2 + \left(\frac{y-\mu_y}{\sigma_y}\right)^2 - 2\rho\frac{(x-\mu_x)(y-\mu_y)}{\sigma_x\sigma_y}$$

where $F(x, y)$ is the quadratic form:

For such r.v.s

$$\begin{aligned} E[X] &= \mu_x \\ E[Y] &= \mu_y \\ \text{Var}(X) &= \sigma_x^2 \\ \text{Var}(Y) &= \sigma_y^2 \\ \text{Cov}(X, Y) &= \rho\sigma_x\sigma_y \end{aligned}$$

where $-1 \leq \rho \leq +1$. (The bivariate normal distribution can be generalized to an arbitrary number of random variables. Such are called jointly Gaussian r.v.s.)

2.7 Gaussian Properties

+ Jointly Gaussian r.v.s X and Y are statistically independent if and only if (iff) they are uncorrelated, i.e., $\rho = 0$.

+ A linear combination of an arbitrary number of jointly Gaussian r.v.s is a Gaussian r.v.

+ **Conditional Gaussian:** Let r.v.s X and Y be jointly Gaussian. Then the conditional pdf of X given $Y = y$ is a single variable Gaussian pdf with

$$\begin{aligned} E(X|Y = y) &= \mu_x + \rho\frac{\sigma_x}{\sigma_y}(y - \mu_y) \\ \text{Var}(X|Y = y) &= \sigma_x^2(1 - \rho^2). \end{aligned}$$

+ **Gaussian Moments:** Let X be a Gaussian random variable with mean zero and variance σ^2 , i.e., $\mathcal{N}(0, \sigma^2)$. Then

$$E(X^{2n}) = 1 \times 3 \times \dots \times (2n - 1)\sigma^{2n}$$

and

$$E(X^{2n-1}) = 0$$

for $n = 1, 2, 3, \dots$

+ Let X and Y be zero mean jointly Gaussian r.v.s with equal variances σ^2 and $\rho = 0$ (i.e., they are statistically independent). Then the derived r.v.s $R = \sqrt{X^2 + Y^2}$ and $\Theta = \arctan(Y/X)$ (four quadrant inverse tangent) are themselves statistically independent and R is Rayleigh with parameter σ and Θ is uniform on $[0, 2\pi)$.

Best estimator is a linear estimator in this case!

