

Monday April 5 - Lect. 32

Maximum Likelihood
Est.

$$\forall \theta \in \Lambda \quad f_{\theta}(y)$$

$\Rightarrow f_{\theta}(\cdot)$ as a prob. density on Γ

$\Rightarrow f_{\theta}(y)$ as a funct. of θ on Λ where y is now fixed.

\hookrightarrow here called the likelihood function

Might slightly change notation

$$f_{\theta}(y) = f(y, \theta) \text{ or } f(y; \theta)$$

Then the max likelihood estimate (MLE)

$$\hat{\theta}_{ML}(y) = \underset{\theta \in \Lambda}{\operatorname{argmax}} f_{\theta}(y)$$

possible: * there is no maximizing θ

* may not be unique.

If $G(\cdot)$ is an increasing funct. on \mathbb{R}

$$\operatorname{argmax}_{\theta} f_{\theta}(y) = \operatorname{argmax}_{\theta} G(f_{\theta}(y))$$

Pick $G(\cdot)$ to simplify ... Most common to use natural log ...

$$\hat{\theta}_{ML}(y) = \operatorname{argmax}_{\theta} \underbrace{\log f_{\theta}(y)}_{\text{log-likelihood function.}}$$

Subject to "smoothness conditions" a necc. cond. that $\hat{\theta}_{ML}(y)$ be an ML estimator is that it is a soln to the likelihood equation

$$\left. \frac{\partial}{\partial \theta} \log f_{\theta}(y) \right|_{\theta = \hat{\theta}_{ML}(y)} = 0.$$

- If θ is vector ... replace $\frac{\partial}{\partial \theta}$ with gradient...
- $\frac{\partial}{\partial \theta} \log f_{\theta}(y)$ is called the score function. The roots have interesting properties.

Example $Y = [Y_1, \dots, Y_n]^T$ iid sample $N(\mu, \sigma^2)$ where $\theta = [\mu, \sigma^2]^T$ is unknown ... ($\mu \in \mathbb{R}, \sigma^2 > 0$)

$$f(y; \mu, \sigma^2) = \frac{1}{(2\pi)^{n/2} \sigma^n} \exp\left\{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2\right\}$$

$$\Rightarrow \log f(y; \mu, \sigma^2) = \text{const} - n \log \sigma - \frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2$$

Take $\frac{\partial}{\partial \mu}$ and $\frac{\partial}{\partial \sigma}$ and set equal to 0

$$\Rightarrow \left. \begin{aligned} \frac{1}{\sigma^2} \sum (y_i - \mu) &= 0 \\ -\frac{n}{\sigma^3} + \frac{1}{\sigma^3} \sum (y_i - \mu)^2 &= 0 \end{aligned} \right\} \Rightarrow \begin{aligned} \hat{\mu} &= \bar{y} = \frac{1}{n} \sum_{i=1}^n y_i \\ \hat{\sigma}^2 &= \frac{1}{n} \sum_{i=1}^n (y_i - \hat{\mu})^2 \end{aligned}$$

Computation of ML

Define $L(\theta, y) = \log f_{\theta}(y)$. ∇_{θ} = gradient.

$$\nabla_{\theta} L(\theta, y) = \left[\frac{\partial}{\partial \theta_1} L(\theta, y), \frac{\partial}{\partial \theta_2} L(\theta, y), \dots, \frac{\partial}{\partial \theta_m} L(\theta, y) \right]^T$$

and similarly ...

$$\nabla_{\theta}^2 L(\theta, y) = \left[\frac{\partial^2 L(\theta, y)}{\partial \theta_i \partial \theta_j} \right]_{i,j=1}^m$$

Want to solve ... $\nabla_{\theta} L(\theta, y) = 0$.

Consider Newton Raphson method for finding these zeros.

Taylor Series ... $\hat{\theta}^{(j)}$ denote an estimate computed @ j th iteration

$$\nabla_{\theta} L(\theta, y) = \nabla_{\theta} L(\hat{\theta}^{(j)}, y) + \nabla_{\theta}^2 L(\hat{\theta}^{(j)}, y) (\theta - \hat{\theta}^{(j)})$$

Ignore HOTS set left hand side to zero and solve for $\theta = \hat{\theta}^{(j+1)}$ + H.O.T.s.

$$0 = \nabla_{\theta} L(\hat{\theta}^{(j)}, y) + \nabla_{\theta}^2 L(\hat{\theta}^{(j)}, y) \cdot$$

$$\Rightarrow \hat{\theta}^{(j+1)}(y) = \hat{\theta}^{(j)}(y) - \left[\nabla_{\theta}^2 L(\hat{\theta}^{(j)}, y) \right]^{-1} \nabla_{\theta} L(\hat{\theta}^{(j)}, y) \cdot (\hat{\theta}^{(j+1)} - \hat{\theta}^{(j)})$$

This gives a seq. $\{\hat{\theta}^{(j)}\}$ which :

- may or may not converge.
- even if it does converge ... the limit could corresp. to

local min
local max
inflection pt.

Fact $L(\theta, y)$ is concave \implies N.R. converges to a global max.

[If $L(\theta, y)$ is twice differentiable then
concave $\iff \nabla_{\theta}^2 L(\theta, y) \leq 0 \quad \forall \theta \in \Delta$]

Possible Simplification

In N.R. must invert max matrix @ every iteration:

- If $\hat{\theta}^{(0)}$ is a good estimat. then might have

$$\nabla_{\theta}^2 L(\hat{\theta}^{(j)}, y) \approx \nabla_{\theta}^2 L(\hat{\theta}^{(0)}, y).$$

so might just use one ...

- $$\nabla_{\theta}^2 L(\hat{\theta}^{(0)}, y) \approx E_{\hat{\theta}^{(0)}} \left\{ \nabla_{\theta}^2 L(\hat{\theta}^{(0)}, Y) \right\}$$

$$= -I_{\hat{\theta}^{(0)}} \quad \text{info. matrix}$$

$$\Rightarrow \hat{\theta}^{(j+1)} = \hat{\theta}^{(j)} + I_{\hat{\theta}^{(0)}}^{-1} (\nabla_{\theta} L(\hat{\theta}^{(j)}, y)).$$

ML Properties

* ML estimates may be biased.

$$\hat{\mu} = \frac{1}{n} \sum y_i$$

$$\hat{\sigma}^2 = \frac{1}{n} \sum (y_i - \hat{\mu})^2$$

$$E_{\theta} \{ \hat{\mu} \} = \mu \quad \forall \theta$$

$$E_{\theta} \{ \hat{\sigma}^2 \} = \frac{n-1}{n} \sigma^2$$

is biased.

* In "regular" sit. where an efficient est. exists it will be an ML estimator