

# Wednesday March 24 Lecture 27

## Suff. Stat. Examples

#1  $Y_k = A + W_k \quad k = 0, 1, \dots, N-1 \quad W_k \sim N(0, \sigma^2) \text{ iid}$   
 $A \in \mathbb{R}$  param.

$$f_{Y;A}(y) = \left(\frac{1}{2\pi\sigma^2}\right)^{N/2} \exp\left\{-\frac{1}{2\sigma^2} \sum_{k=0}^{N-1} (y_k - A)^2\right\}$$

exp and ...

$$= \underbrace{\left(\frac{1}{2\pi\sigma^2}\right)^{N/2} \exp\left\{-\frac{1}{2\sigma^2} \left(NA^2 - 2A \sum_{k=0}^{N-1} y_k\right)\right\}}_{g(T(y); A)} \underbrace{\exp\left\{-\frac{1}{2\sigma^2} \sum_{k=0}^{N-1} y_k^2\right\}}_{h(y)}$$

Then N.P. Fact. Thm  $\Rightarrow T(y) = \sum_{k=0}^{N-1} y_k$  is suff for A.

$\Rightarrow$  Actually any one-to-one function of  $T(y)$  would do.

(#2) Same model ...  $A=0$ ,  $\sigma^2$  unknown param.

$$f_{Y; \sigma^2}(y) = \underbrace{\frac{1}{(2\pi\sigma^2)^{N/2}} \exp\left\{-\frac{1}{2\sigma^2} \sum_{k=0}^{N-1} y_k^2\right\}}_{g(T(y); \sigma^2)} \cdot \underbrace{1}_{h(y)}$$

$$\Rightarrow T(y) = \sum_{k=0}^{N-1} y_k^2 \text{ is suff. for } \sigma^2.$$

Another approach: Go to Definition

$Y, T(Y)$  rvs. pdfs from model  $f_{Y; \theta}(y)$

$$f_{Y, T; \theta}(y, t) = f_{Y; \theta}(y) \delta(T(y) = t)$$

Then define

$$f_{Y|T; \theta}(y|t) = \frac{f_{Y, T; \theta}(y, t)}{f_{T; \theta}(t)} = \frac{f_{Y; \theta}(y) \delta(T(y) = t)}{f_{T; \theta}(t)}$$

Suff. if this is  
indep of  $\theta$ .

Now for (2)  $T(y) = \sum_{k=0}^{N-1} y_k^2 = \sigma^2 \left\{ \begin{array}{l} \text{a chi-sq. rv.} \\ \text{with } N \text{ d.o.f.} \end{array} \right\}$

If  $S$  is  $\chi^2$   $N$  d.o.f.

$$f_S(s) = \begin{cases} \frac{1}{2^{N/2} \Gamma(N/2)} e^{-s/2} s^{N/2-1} & s > 0 \\ 0 & s < 0 \end{cases}$$

$$\frac{1}{\sigma^2} = S \implies f_{T; \sigma^2}(t) = \begin{cases} \frac{1}{\sigma^2} \frac{1}{2^{N/2} \Gamma(N/2)} e^{-t/2\sigma^2} \left(\frac{t}{\sigma^2}\right)^{N/2-1} & t > 0 \\ 0 & t < 0. \end{cases}$$

$$\frac{1}{\sigma^2} f_{Y|T; \theta}(y|t) = \frac{\frac{1}{(2\pi\sigma^2)^{N/2}} \exp\left\{-\frac{1}{2\sigma^2} t(y)\right\} \delta(T(y)=t)}{\frac{1}{\sigma^2} \frac{1}{2^{\frac{N}{2}} \Gamma(\frac{N}{2})} e^{-t/2\sigma^2} \left(\frac{t}{\sigma^2}\right)^{\frac{N}{2}-1}}$$

$$= \frac{\Gamma(\frac{N}{2})}{\pi^{N/2} t^{\frac{N}{2}-1}} \delta(T(y)=t).$$

This does not dep on  $\sigma^2$

$\therefore T$  is suff. for  $\sigma^2$ .

### ③ Phase of Sinusoid

$$Y_n = A \cos(2\pi f_0 n + \phi) + W_n \quad n=0, 1, \dots, N-1$$

$\hookrightarrow \sim N(0, \sigma^2)$  iid.

$A, f_0, \sigma^2$  all known.

Parameter is  $\phi$ .

$$f_{Y; \phi}(y) = \frac{1}{(2\pi\sigma^2)^{N/2}} \exp\left\{-\frac{1}{2\sigma^2} \sum_{n=0}^{N-1} (y_n - A \cos(2\pi f_0 n + \phi))^2\right\}$$

$$\sum y_n^2 - 2A \underbrace{\left(\sum y_n \cos 2\pi f_0 n\right)}_{T_1(y)} \cos \phi + 2A \underbrace{\left(\sum y_n \sin 2\pi f_0 n\right)}_{T_2(y)} \sin \phi + \sum A^2 \cos^2(2\pi f_0 n + \phi).$$

$$\Rightarrow f_{Y; \phi}(y) = g(T_1(y), T_2(y); \phi) \cdot h(y)$$

Example  $Y_n = r^n + W_n$   $W_n \sim N(0, \sigma^2)$  iid  
 $n = 0, 1, \dots, N-1$  ↘ Known.

Param.  $r$ .

$$f_{Y;r}(y) = \frac{1}{(2\pi\sigma^2)^{N/2}} \exp \left\{ -\frac{1}{2\sigma^2} \sum_{n=0}^{N-1} (y_n - r^n)^2 \right\}$$

$$\sum_{n=0}^{N-1} y_n^2 - 2 \sum_{n=0}^{N-1} r^n y_n + \sum_{n=0}^{N-1} r^{2n}$$

(The middle term  $- 2 \sum_{n=0}^{N-1} r^n y_n$  is circled in blue in the original image.)

Cannot factor into a funct.  
of the  $y_n$   $n=0, \dots, N-1$  alone  
(at least a funct indep of  $r$ ).

## Using Sufficiency to MVUE

(#1)  $Y_k = A + W_k$      $W_k \sim N(0, \sigma^2)$  iid     $A \in \mathbb{R}$  param.

$T(y) = \sum_{k=0}^{N-1} y_k$  is suff. for  $A$ .

↑  
known

### Two approaches

(a) Find any unbiased est. of  $A$  say  $\tilde{A}(y) = y_0$

$\Rightarrow \hat{A}(y) = E\{\tilde{A} | T(y)\}$

(b) Just find a funct.  $g(\cdot)$  st.

$\hat{A}(y) = g(T(y))$   
is unbiased.

Approach (a)  $\hat{A} = E\{\tilde{A} | T\}$   $\tilde{A}$  and  $T$  are jointly Gaussian with means

$$E\{\tilde{A}\} = E\{Y_0\} = A$$

$$E\{T\} = NA$$

and some cov. matrix

$$\Sigma = \begin{bmatrix} \text{Var}\{\tilde{A}\} & \text{Cov}\{\tilde{A}, T\} \\ \text{Cov}\{T, \tilde{A}\} & \text{Var}\{T\} \end{bmatrix}$$

1st day ...

$$E\{\tilde{A} | T=t\} = \underbrace{E\{\tilde{A}\}}_A + \frac{\text{Cov}\{\tilde{A}, T\}}{\underbrace{\text{Var}\{T\}}_{N\sigma^2}} \left[ t - \underbrace{E\{T\}}_{NA} \right]$$

$$\text{Cov}\{\tilde{A}, T\} = \text{Cov}\left\{Y_0, \sum_{k=0}^{N-1} Y_k\right\} = \sum_{k=0}^{N-1} \text{Cov}\{Y_0, Y_k\} = \sigma^2$$

$$E\{\tilde{A} | T=t\} = A + \frac{\sigma^2}{N\sigma^2} (t - NA)$$

$$= \frac{1}{N} t = \frac{1}{N} \sum_{k=0}^{N-1} y_k$$

Approach is easier.