

EE538 DSPI

Module 24

Outline:

- Nonparametric estimation of autocorrelation and power spectrum
 - Sect. 12.1, 12.2.3
- Sum of sinewaves model - Sect. 12.5.2
- DT random processes - Appendix A

- Parametric Spectral Estimation
- with emphasis on harmonic retrieval methods
- i.e., estimating frequencies and amplitudes of superimposed sinewaves (or geometric sequences) embedded in noise

- parametric = model-based
- Why parametric?
- to achieve higher resolution of sharp spectral peaks than that achieved with nonparametric spectral estimation
- the latter is primarily Wiener-Kinchen Theorem

- power or energy density spectrum of a DT signal $x[n]$ is

$$S_{xx}(\omega) = |X(\omega)|^2$$

- where: $x[n] \xleftrightarrow{\text{DTFT}} X(\omega)$

- $S_{xx}(\omega) = \sum_{m=-\infty}^{\infty} r_{xx}[m] e^{-j m \omega}$

- where: $r_{xx}[m] = \sum_{n=-\infty}^{\infty} x[n] x^*[n-m]$

• Stochastic Case:

$$S_{xx}(\omega) = E\{|X(\omega)|^2\}$$

• $r_{xx}[m] \xleftrightarrow{\text{DTFT}} S_{xx}(\omega)$

• where:

$$r_{xx}[m] = E\{x[n]x^*[n-m]\}$$

• where: $E\{\cdot\}$ = expectation operator

• $x[n]$ assumed to be wide-sense stationary - see Appendix A

• in practical case, given N samples of a wide sense stationary random process $x[n]$, $n=0, 1, \dots, N-1$

• autocorrelation is estimated as

$$\hat{r}_{xx}[m] = \frac{1}{N-m} \sum_{n=0}^{N-m-1} x[n] x^*[n-m]$$

for $m=0, 1, \dots, N-1$

- note: upper limit in sum
- for a given m , there are $N-m$ pairs of samples separated by m units of DT to average over in order to:
- achieve uncorrelatedness/independence amongst different constituent signals comprising $x[n]$
- average out the effects of noise

Case Study: Sum of Sinewaves
(SoS) in Noise

$$x[n] = \sum_{i=1}^P A_i e^{j(\omega_i n + \phi_i)} + v[n]$$

- A_i, ω_i : deterministic but unknown
- ϕ_i : independent and identically (i.i.d.) uniform random variables over $[0, 2\pi)$

$v[n]$: contribution from noise

• assumptions :

i.) zero mean : $E\{v[n]\} = 0$

ii.) uncorrelated / independent
with signal

$E\{v[n] s^*[n-m]\} = 0$ for all m

• where:

$$s[n] = \sum_{i=1}^P A_i e^{j\phi_i} e^{j\omega_i n}$$

iii.) $v[n]$ is a DT WSS
random process

$$r_{vv}[m] = E\{v[n]v^*[n-m]\}$$

• typically assume white noise:

$$r_{vv}[m] = \sigma_w^2 \delta[m]$$

• $E\{\cdot\}$ is a linear operator:

$$E\left\{\sum_{i=1}^Q a_i X_i\right\} = \sum_{i=1}^Q a_i E\{X_i\}$$

- "true" autocorrelation for
SoS in noise:

$$\begin{aligned}
 r_{xx}[m] &= E \left\{ x[n] x^*[n-m] \right\} \\
 &= E \left\{ \sum_{i=1}^P A_i e^{j\phi_i} e^{j\omega_i n} \sum_{l=1}^P A_l e^{-j\phi_l} e^{-j\omega_l (n-m)} \right\} \\
 &= \sum_{i=1}^P \sum_{l=1}^P A_i A_l e^{j\omega_i n} e^{-j\omega_l (n-m)} \\
 &\quad \cdot \underbrace{E \left\{ e^{j\phi_i} e^{-j\phi_l} \right\}}_{\delta_{i-l}} +
 \end{aligned}$$

$$\begin{aligned}
 &+ E\{v[n] s^*[n-m]\} + E\{s[n] v^*[n-m]\} \\
 &+ E\{v[n] v^*[n-m]\}
 \end{aligned}$$

• due to i.i.d. assumption,
for $i \neq l$:

$$E\{e^{j\phi_i} e^{-j\phi_l}\} = E\{e^{j\phi_i}\} E\{e^{-j\phi_l}\}$$

$$\begin{aligned}
 E\{e^{j\phi_i}\} &= \int_0^{2\pi} \frac{1}{2\pi} e^{j\phi} d\phi \\
 &= \frac{1}{j} \frac{1}{2\pi} e^{j\phi} \Big|_0^{2\pi} = \frac{1}{2\pi j} \{1 - 1\} = 0
 \end{aligned}$$

• thus:

$$r_{xx}[m] = \sum_{i=1}^P P_i e^{j\omega_i m} + \sigma_w^2 \delta[m]$$

$P_i = A_i^2, i=1, \dots, P$

for all m

• Compare to time-averaged estimate of autocorrelation

• substitute SoS in noise model into

$$\hat{r}_{xx}[m] = \sum_{n=0}^{N-m-1} x[n] x^*[n-m] \frac{1}{N-m}$$

$$\begin{aligned}
 \hat{r}[m] = & \sum_{i=1}^P \sum_{l=1}^P A_i A_l e^{j(\phi_i - \phi_l)} e^{j\omega_l m} \\
 & \cdot \frac{1}{N-m} \sum_{n=0}^{N-m-1} e^{j(\omega_i - \omega_l)n} \\
 & + \frac{1}{N-m} \sum_{n=0}^{N-m-1} s[n] v^*[n-m] + \\
 & \frac{1}{N-m} \sum_{n=0}^{N-m-1} v[n] s^*[n-m] + \frac{1}{N-m} \sum_{n=0}^{N-m-1} v[n] v^*[n-m]
 \end{aligned}$$

• for lag value m , need to avg.
over enough DT sample pairs
separated by m so that:

$$\frac{1}{N-m} \sum_{n=0}^{N-m-1} e^{j(\omega_1 - \omega_2)n}$$
$$= \frac{1}{N-m} \cdot \frac{1 - e^{j(\omega_1 - \omega_2)(N-m)}}{1 - e^{j(\omega_1 - \omega_2)}}$$

$$= \frac{1}{N-m} \frac{\sin \left[\frac{N-m}{2} (\omega_i - \omega_a) \right]}{\sin \left[\frac{1}{2} (\omega_i - \omega_a) \right]}$$

$$\lim_{N \rightarrow \infty} \underbrace{\frac{\sin \left[\frac{N-m}{2} (\omega_i - \omega_a) \right]}{\sin \left[\frac{1}{2} (\omega_i - \omega_a) \right]}}_{= \delta[i-a]}$$

for fixed $m < \infty$

$$= \frac{1}{\sin \left[\frac{1}{2} (\omega_i - \omega_a) \right]} \cdot \frac{\sin \left[\frac{N-m}{2} (\omega_i - \omega_a) \right]}{N-m}$$

- as lag value m increases,
less number of pairs to
average over
- estimates of $r_{xx}[m]$ become
increasingly unreliable as
 m increases

- general form of nonparametric spectral estimator:

$$\hat{S}_{xx}(\omega) = \sum_{m=-(N-1)}^{N-1} \hat{r}_{xx}[m] W[m] e^{-j\omega m}$$

- where $W[m]$ is a real-valued, symmetric ($W[-m] = W[m]$), tapered ($W[m+1] \leq W[m]$), and $W(\omega) = \text{DTFT}\{W[m]\} > 0$ for all ω

• e.g. Bartlett (Triangle) Window

$$\cdot w[m] = \frac{N - |m|}{N}, \quad m = 0, 1, \dots, M$$

where: $M \leq N-1$

= 0 otherwise

• Expected value of nonparametric spectral estimator (See Sect. 12.2.3)

$$\cdot E\{\hat{S}_{xx}(f)\} = \sum_{m=-(N-1)}^{N-1} E\{\hat{r}_{xx}[m]\} w[m] e^{-j\omega m}$$

$$= \sum_{m=-\infty}^{\infty} r_{xx}[m] w[m] e^{-j\omega m}$$

$m=-\infty$

$$= S_{xx}(\omega) \circledast W(\omega)$$

↑
true spectrum

↑ periodic convolution

see pg. 303 of P&M text

- convolution operation with $W(\omega)$ causes "smearing" of fine spectral details and "ringing" - noisy background that can mask weak signal components

- Return to SOS in noise Case Study

$$S_{xx}(f) = \text{DTFT}\{r_{xx}[m]\}$$

$$= \sum_{i=1}^P P_i 2\pi \delta_a(\omega - \omega_i) + \sigma_w^2$$

↑
Dirac Delta function for $|\omega| < \pi$

$$E\{\hat{S}_{xx}(\omega)\} = \sum_{i=1}^P P_i W(\omega - \omega_i) + \sigma_w^2 W[0]$$

• Example: $P=3$

