

11/8/2022

Exam 2

- No mgf/char. fn. questions
 - Explain your work!
 - Integrate ^{constants,} polynomials, e^{ax} ; leave more complicated integrals
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Notation for convergence in distribution: $X_n \xrightarrow{d} X$

Note: It is possible for a sequence X_1, X_2, \dots to converge in distr., but to have density fns that do not converge $\forall x \in \mathbb{R}$.

Example: Consider a sequence of rvs X_1, X_2, \dots with

$$f_{X_n}(x) = \begin{cases} 1 + \cos 2\pi nx, & 0 \leq x \leq 1 \\ 0 & \text{elsewhere} \end{cases}$$

$n = 1, 2, \dots$

Then $F_{X_n}(x) \Rightarrow \begin{cases} 0 & x < 0 \\ x & 0 \leq x \leq 1 \\ 1 & x > 1 \end{cases}$ ②

as $n \rightarrow \infty$

So X_n converges in distr. to a unif $[0,1]$ rv

But f_{X_n} does not converge except at $x=0, x=1$



$$e \Rightarrow ae \Rightarrow p \Rightarrow d$$

MS \nearrow

Proof outside scope of class

The Cauchy Criterion for MS convergence

It is possible to show MS convergence without knowing the

limiting rv X . First consider (3)
a sequence x_1, x_2, \dots , where
 $x_i \in \mathbb{R} \forall i$. The Cauchy criterion
states that if $|x_{n+m} - x_n| \rightarrow 0$
as $n \rightarrow \infty$, for any $m \geq 1$, then
 x_n converges to some limit
 x . The converse is also
true.

Proof omitted. (from real
analysis)

For a sequence of rvs X_1, X_2, \dots ,
the Cauchy criterion for
convergence states that

X_1, X_2, \dots converges in MS iff

$$E[|X_{n+m} - X_n|^2] \rightarrow 0 \text{ as } n \rightarrow \infty, \\ \forall m \geq 1.$$

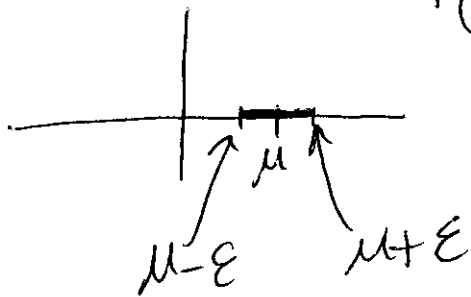
The Chebyshev Inequality

Let X be a rv with finite
mean μ and variance σ^2 . Then

$\forall \varepsilon > 0,$

(4)

$$P(|X - \mu| \geq \varepsilon) \leq \frac{\sigma^2}{\varepsilon^2}$$



Proof. Let $g_1(x) = I_{\{x \in \mathbb{R} : |x - \mu| \geq \varepsilon\}}(x)$

and $g_2(x) = \frac{(x - \mu)^2}{\varepsilon^2}$

for some arbitrary $\varepsilon > 0,$



Now let $\phi(x) = g_2(x) - g_1(x)$
 $\geq 0 \quad \forall x$

So $E[\phi(X)] \geq 0.$

But $E[\phi(X)] = E[g_2(X)] - E[g_1(X)]$

$$= \frac{\sigma^2}{\varepsilon^2} - P(|X - \mu| \geq \varepsilon)$$

since $E[g_1(X)] = 0 \cdot P(|X - \mu| < \varepsilon) + 1 \cdot P(|X - \mu| \geq \varepsilon)$

$$\text{So } \frac{\sigma^2}{\varepsilon^2} - P(|X - \mu| \geq \varepsilon) \geq 0, \quad (5)$$

which means $P(|X - \mu| \geq \varepsilon) \leq \frac{\sigma^2}{\varepsilon^2}$

The Laws of Large Numbers

Consider a sequence of rvs X_1, X_2, \dots , and the sequence of sample means

$$Y_n = \frac{1}{n} \sum_{k=1}^n X_k$$

The Weak Law of Large Numbers (WLLN) states that

$$P(|Y_n - \mu| \geq \varepsilon) \rightarrow 0 \quad \text{as } n \rightarrow \infty$$

if X_1, X_2, \dots is an iid sequence with finite mean μ and variance σ^2 . In other words

$$Y_n \rightarrow \mu \quad \text{in prob.}$$

Proof. First note that

(6)

$$E[Y_n] = \mu, \text{ and}$$

$$\text{Var}(Y_n) = \frac{\sigma^2}{n}$$

then, from the CI,

$$P(|Y_n - \mu| \geq \varepsilon) \leq \frac{\text{Var}(Y_n)}{\varepsilon^2} = \frac{\sigma^2}{n\varepsilon^2}$$

$\rightarrow 0$ as $n \rightarrow \infty$,
 $\forall \varepsilon > 0$.