

Notes presented on
chalkboard 11/15/2022

Theorem. (Strong Law of Large Numbers)

Let X_n be a sequence of iid
rvs with mean μ and variance
 σ^2 . Then

if the mean and variance are finite,

$$Y_n = \frac{1}{n} \sum_{k=1}^n X_k \xrightarrow{\text{a.e.}} \mu \quad \text{as } n \rightarrow \infty$$

Proof. Omitted.

So for the example

$$X_k = s + W_k$$

if the W_k 's are independent and have the same variance σ^2 , then the sample mean Y_n converges to s with probability 1.

But the laws of large numbers tell us something more fundamental about our axiomatic-based prob. measure. Consider a rv X and a set $A \subset B(\mathbb{R})$. Let X_1, X_2, \dots be an iid sequence of

r.v.s with the same distributions as X . Now let

$$Y_k = \begin{cases} 1 & \text{if } X_k \in A \\ 0 & \text{if } X_k \notin A \end{cases}$$

$$\text{Then } E[Y_k] = P(X_k \in A) = P(X \in A)$$

The Y_k 's are iid with mean $P(X \in A)$, and $\frac{1}{n} \sum_{k=1}^n Y_k$ is the relative frequency of the event $\{X \in A\}$. The SLLN tells us that

$$\frac{1}{n} \sum_{k=1}^n Y_k \rightarrow P(X \in A) \text{ w.p.1, as } n \rightarrow \infty$$

So with prob. 1, the relative frequency approach to prob. converges to the axiom-based $P(X \in A)$.

The Central Limit Theorem

Let X_n be a sequence of iid rvs with finite mean μ and finite variance σ^2 . Then if

$$Z_n = \frac{\sum_{j=1}^n X_j - n\mu}{\sigma\sqrt{n}}$$

then $Z_n \xrightarrow{d} Z$ as $n \rightarrow \infty$
where Z is $N(0,1)$

i.e., $F_{Z_n}(z) \rightarrow \Phi(z) = \int_{-\infty}^z \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx$

Proof. Will use without proof the following:

Lemma. Let Z_n be a sequence of rvs with cdfs F_{Z_n} and mgfs $\Phi_{Z_n}(s)$, and let Z be a rv with cdf F_Z and mgf Φ_Z . Then if $\Phi_{Z_n}(s) \rightarrow \Phi_Z(s)$

$\forall s \in \mathbb{C}$, then $F_{z_n}(z) \rightarrow F_z(z) \forall z \in \mathbb{R}$.

So it is sufficient to show that $\Phi_{Z_n}(s) \rightarrow e^{s^2/2}$

Consider $\Phi_X\left(\frac{s}{\sqrt{n}}\right) = E\left[e^{sX_k/\sqrt{n}}\right] \forall k$

Now considering the case $\mu=0, \sigma^2=1$, we have

$$\begin{aligned}\Phi_{Z_n}(s) &= E\left[e^{s \sum_{k=1}^n X_k/\sqrt{n}}\right] \\ &= E\left[\prod_{k=1}^n e^{sX_k/\sqrt{n}}\right] = \prod_{k=1}^n E\left[e^{sX_k/\sqrt{n}}\right] \\ &= \prod_{k=1}^n \Phi_X\left(\frac{s}{\sqrt{n}}\right) = \left(\Phi_X\left(\frac{s}{\sqrt{n}}\right)\right)^n\end{aligned}$$

Let $L(s) = \log \Phi_X(s)$. Then to show $\Phi_{Z_n}(s) \rightarrow e^{s^2/2}$, it is sufficient to show $nL\left(\frac{s}{\sqrt{n}}\right) = \log\left[\left(\Phi_X\left(\frac{s}{\sqrt{n}}\right)\right)^n\right] \rightarrow \frac{s^2}{2}$

Will need:

$$L(0) = 0$$

$$L'(0) = \mu = 0$$

$$L''(0) = E[X^2] = 1$$

Now

$$\lim_{n \rightarrow \infty} nL\left(\frac{s}{\sqrt{n}}\right) = \lim_{n \rightarrow \infty} \frac{L\left(\frac{s}{\sqrt{n}}\right)}{n^{-1}}$$

$$= \lim_{n \rightarrow \infty} \frac{-L'\left(\frac{s}{\sqrt{n}}\right) n^{-3/2} s}{-2n^{-2}} \quad (\text{l'Hopital})$$

$$= \lim_{n \rightarrow \infty} \frac{L'\left(\frac{s}{\sqrt{n}}\right) s}{2n^{-1/2}}$$

$$= \lim_{n \rightarrow \infty} \frac{-L''\left(\frac{s}{\sqrt{n}}\right) n^{-3/2} s^2}{-2n^{-3/2}} \quad (\text{l'Hopital again})$$

$$= \lim_{n \rightarrow \infty} \frac{L''\left(\frac{s}{\sqrt{n}}\right) s^2}{2} = \frac{s^2}{2}$$

For general μ, σ^2 use result
for $\mu=0, \sigma^2=1$ with $Y_k = \frac{X_k - \mu}{\sigma}$