

EE 641 Midterm Exam
October 18, Fall 2013

Name: Key

Instructions

The following is an in-class closed-book exam.

- This exam contains 3 problems worth a total of 108 points.
- You may not use any notes, textbooks, or calculators.
- You are allowed 50 minutes to complete the exam. Hand in the exam after that period **whether or not you have completed it.**

Good luck.

Problem 1. (36pt)

Let $X \sim N(0, R)$ where R is a $p \times p$ symmetric positive-definite matrix with an eigen decomposition of the form $R = E\Lambda E^t$.

a) Calculate the covariance of $\tilde{X} = E^t X$, and show that the components of $\tilde{X}$ are jointly independent Gaussian random variables.

b) Show that if $Y = E\Lambda^{1/2}W$ where $W \sim N(0, I)$, then $Y \sim N(0, R)$.

c) How can this result be of practical value?

$$\begin{aligned} \text{a)} \quad E[\tilde{X}\tilde{X}^t] &= E[E^t X X^t E] \\ &= E^t E[X X^t] E \\ &= E^t E \Lambda E^t E \\ &= \Lambda \end{aligned}$$

$$\begin{aligned} \text{b)} \quad Y &= E \Lambda^{1/2} W \\ \mu_Y &= E[Y] = E[E \Lambda^{1/2} W] = E \Lambda^{1/2} E[W] \\ R_Y &= E[E \Lambda^{1/2} W W^t \Lambda^{1/2} E^t] = 0 \\ &= E \Lambda^{1/2} E[W W^t] \Lambda^{1/2} E^t \\ &= E \Lambda^{1/2} I \Lambda^{1/2} E^t \\ &= E \Lambda E^t = R \end{aligned}$$

Also, Y must be jointly Gaussian because it is a linear transform of a Gaussian random vector W .

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b) continued...

Another solution

$$A = E \Lambda^{1/2} \quad Y = AX$$

$$A^{-1} = \Lambda^{-1/2} E$$

$$(A^{-1})^T = E^T \Lambda^{-1/2}$$

$$p_Y(y) = \frac{1}{|A|} p_W(A^{-1}y) = \frac{1}{2} \exp\left\{-\frac{1}{2} (A^{-1}y)^T I (A^{-1}y)\right\}$$

$$= \frac{1}{2} \exp\left\{-\frac{1}{2} y^T E \Lambda^{-1/2} I \Lambda^{-1/2} E^T y\right\}$$

$$= \frac{1}{2} \exp\left\{-\frac{1}{2} y^T E \Lambda^{-1} E^T y\right\}$$

$$= \frac{1}{(2\pi)^{p/2}} |R|^{-1/2} \exp\left\{-\frac{1}{2} y^T R^{-1} y\right\}$$

$$R^{-1} = E \Lambda^{-1} E^T \quad R = E \Lambda E^T$$

$$\Rightarrow Y \sim N(0, R)$$

c) Part b) provides a practical method for generating Gaussian random vectors with any desired covariance.

Problem 2. (36pt)

Let X_s be a zero-mean 2-D Gaussian AR process indexed by $s = (s_1, s_2)$, and let the MMSE causal prediction filter be given by

$$h_s = \rho \delta_{s_1-1, s_2} + \rho \delta_{s_1, s_2-1}$$

and the causal prediction variance be given by σ_c^2 . Determine (σ_{NC}^2, g_s) the noncausal prediction variance and the noncausal prediction filter.

$$\frac{\sigma_{nc}^2}{1-G(\omega)} = \frac{\sigma_c^2}{|1-H(\omega)|^2} \Rightarrow$$

$$\sigma_{nc}^2 (\delta_n - h(n)) * (\delta_n - h(-n)) = \sigma_c^2 (\delta_n - g(n))$$

$$(\delta_s - h(s)) * (\delta_s - h(-s)) = \frac{\sigma_c^2}{\sigma_{nc}^2} (\delta_s - g(s))$$

$\begin{matrix} & -\rho & 1 & \\ -\rho & & & \end{matrix}$ auto correlate $\Rightarrow$ $\begin{matrix} & 0 & -\rho & \rho^2 \\ -\rho & 1+2\rho^2 & -\rho & \\ \rho^2 & -\rho & 0 & \end{matrix}$

$\sigma_{nc}^2 = \frac{1}{1+2\rho^2}$

$$\Rightarrow \begin{bmatrix} 0 & \frac{-\rho}{1+2\rho^2} & \frac{\rho^2}{1+2\rho^2} \\ \frac{-\rho}{1+2\rho^2} & 0 & \frac{-\rho}{1+2\rho^2} \\ \frac{\rho^2}{1+2\rho^2} & \frac{-\rho}{1+2\rho^2} & 0 \end{bmatrix} = -g(s)$$

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$$g(s) = \begin{cases} 0 & s = (0, 0) \\ \frac{-p}{1+2p^2} & s = (1, 0), (-1, 0), (0, 1), (0, -1) \\ \frac{p^2}{1+2p^2} & s = (1, -1), (-1, 1) \\ 0 & \text{otherwise} \end{cases}$$

Problem 3. (36pt)

Show that if X is a zero-mean GMRF with inverse covariance, B , then $B_{s,r} = 0$ if and only if x_s is conditionally independent of X_r given $\{X_i : \text{for } i \neq s, r\}$.

let $B = \begin{bmatrix} \frac{1}{\sigma^2} & \beta^* \\ A & C \end{bmatrix}$ $x = \begin{bmatrix} x_s \\ y \end{bmatrix}$
w.l.o.g.

Then

$$p(x_s | y) = \frac{1}{z} \exp \left\{ -\frac{1}{2} x^* B x \right\}$$

$$\text{where } z = \int_{\mathbb{R}} p(x_s, y) dx_s$$

$$x^* B x = \frac{1}{\sigma^2} x_s^2 + 2 \beta^* y x_s + \text{constant}$$

$$= \frac{1}{\sigma^2} (x_s - \sigma^2 \beta^* y)^2 + \text{constant}$$

$$p(x_s | y) = \frac{1}{z} \exp \left\{ -\frac{1}{2\sigma^2} (x_s - \sigma^2 \beta^* y)^2 \right\}$$

$$= \frac{1}{\sqrt{2\pi}\sigma^2} \exp \left\{ -\frac{1}{2\sigma^2} (x_s - \sigma^2 \beta^* y)^2 \right\}$$

$$\sim N(\sigma^2 \beta^* y, \sigma^2)$$

$$E[x_s | y] = \sigma^2 \beta^* y \quad \text{Var}[x_s | y] = \sigma^2$$

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Since B is positive definite $\Rightarrow \sigma^2 > 0$

so

X_s is conditionally independent
of X_i given $\{X_i \text{ for } i \neq s, n\}$

$\Leftrightarrow p(x_s | x_n^{i \neq s})$ is not a function
of x_n

$\Leftrightarrow \sigma^2 \beta^T y$ is not a function of
 x_n

$\Leftrightarrow \beta_1 = 0$

$\Leftrightarrow B_{sn} = 0 = B_{ns}$