

Lecture 20

Application: Segmentation

X - Class labels

Y - Image Pixels

$$X_s = \begin{cases} 1 & \rightarrow \text{object} \\ 0 & \rightarrow \text{background} \end{cases}$$

$$p(Y|X) = \prod_{s \in S} p(Y_s|X_s)$$

$p(Y_s|0) \leftarrow$ distribution of background
 $p(Y_s|1) \leftarrow$ distribution of object

$X \leftarrow$ using model MRF

$$\log p(Y, X) = \log p(Y|X) + \log p(X)$$

$$\log p(Y|X) = \sum_{s \in S} \log p(Y_s|X_s)$$

$$\begin{aligned} \log p(X) &= \log \left\{ \frac{1}{Z} \exp \left\{ - \sum_{\{i, j\} \in C} \beta \psi(x_i, x_j) \right\} \right\} \\ &= - \sum_{\{i, j\} \in C} \beta \psi(x_i, x_j) + C \end{aligned}$$

Define

$$l(y_s | x_s) \stackrel{\Delta}{=} -\log p(y_s | x_s)$$

$$\log p(x, y) = -\sum_{s \in S} l(y_s | x_s) - \sum_{\{i, j\} \in C} \beta t(x_i, x_j)$$

$$\hat{x}_{MAP} = \underset{x}{\operatorname{argmin}} \left\{ \sum_{s \in S} l(y_s | x_s) + \sum_{\{i, j\} \in C} \beta t(x_i, x_j) \right\}$$

Find the segmentation \hat{x}_{MAP} which

- 1) Fits data
 - 2) Minimizes boundary length
-

Optimization

$$\text{Let } u(x) = \sum_{s \in S} l(y_s | x_s) + \sum_{\{i, j\} \in C} \beta_{ij} \phi(x_i, x_j)$$

By Bayes rule

$$p(x | y) = \frac{1}{Z} \exp\{-u(x)\}$$

$$p(x_s | x_i, i \neq s, y) = \frac{\exp\{-u(x_s, x_i, i \neq s)\}}{\sum_{K=1}^N \exp\{-u(x_s = K, x_i, i \neq s)\}}$$

The conditional "mode" for x_s is

$$\begin{aligned} \hat{x}_s &= \underset{x_s}{\operatorname{argmax}} p(x_s | x_i, i \neq s, y) \\ &= \underset{x_s}{\operatorname{argmax}} \exp\{-u(x_s, x_i, i \neq s)\} \\ &= \underset{x_s}{\operatorname{argmax}} -u(x_s, x_i, i \neq s) \end{aligned}$$

This is coordinate ascent / Gauss-Seidel at location s .

Hence the name \circledast Iterated Conditional Modes (ICM)

ICM: iteratively apply replacement rule.

$$X_s^{(k+1)} = \operatorname{argmin}_{x_s} \{ U(x_s, X_{i \neq s}) \}$$

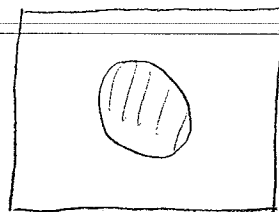
$$X_s^{(k+1)} = \operatorname{argmin}_K \left\{ \ell(y_s | K) + \beta \sum_{i \in \mathcal{S}} \#(K, X_i^{(k)}) \right\}$$

$$= \operatorname{argmin}_K \left\{ \ell(y_s | K) + \beta V(K, X_{\mathcal{S}}) \right\}$$

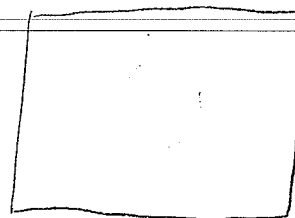
Properties:

1) Convergence \rightarrow yes

2) Global Convergence \rightarrow no



X^a



X^b

$$U(X^b) < U(X^a)$$

but local operation will not get you from X^a to X^b

Lecture 21

Simulated Annealing

Consider the MRF X with distribution

$$p(\hat{x}) = \frac{1}{Z} e^{-\frac{1}{T} u(\hat{x})}$$

If

1) $T = 1 \Rightarrow$ posterior distribution of X given Y

2) $T = \infty \Rightarrow$ uniform on $[1, 2]^N$

3) $T \rightarrow 0 \Rightarrow$ distribution concentrated around X

$$\hat{X}_{MAP} = \underset{X}{\operatorname{argmin}} u(X)$$

First idea

1) pick $T = .0001$

2) generate Markov chain of X^i 's using Gibbs sampler

3) wait for stationary behavior

4) pick any sample segmentation

Bad idea: as $T \rightarrow 0$ 3) takes longer

Second idea (Simulated annealing)

- 1) Choose an annealing schedule
 T_n for iteration n
 $T_{n+1} < T_n$
- 2) Use Gibbs Sampler algorithm with temperature T_n at iteration n
- 3) Wait until T_n get small and pick a sample segmentation

Replacement distribution

$$\begin{aligned} W_s &\sim \frac{\exp\left\{-\frac{1}{T} U(x_s, x_i \text{ } i \neq s)\right\}}{\sum_{k=1}^M \exp\left\{-\frac{1}{T} U(x_s=k, x_i \text{ } i \neq s)\right\}} \\ &= \frac{\exp\left\{-\frac{1}{T} (\ell(y_s | x_s) + \beta V(x_s, x_{2s}))\right\}}{\sum_{k=1}^M \exp\left\{-\frac{1}{T} (\ell(y_s | k) + \beta V(k, x_{2s}))\right\}} \\ &= p(x_s | x_i \text{ } i \neq s) \end{aligned}$$

if $T \rightarrow 0 \Rightarrow \text{ICM}!$

Optimal
Annealing Schedule

N - # of pixels

$$\Delta = \operatorname{argmax}_x U(x) - \operatorname{argmin}_x U(x)$$

K - # of pixels replaced + 1

$$T(K) = \frac{N\Delta}{\log K}$$

Example

if $\Delta = 1$ $N = 250,000$

$$\frac{N\Delta}{\log K} = \frac{1}{2}$$

$$K = e^{500,000}$$

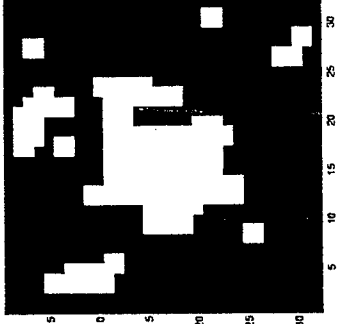
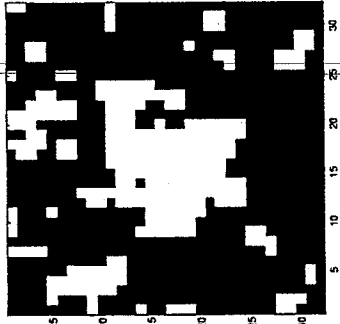
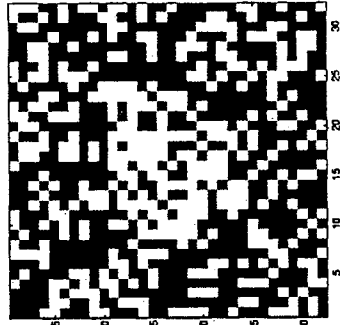
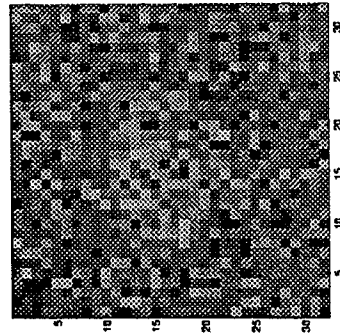
"a long time"

Typical annealing schedule

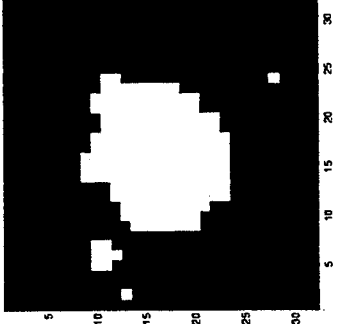
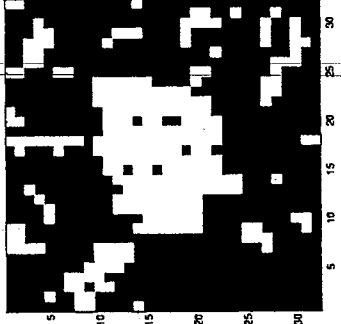
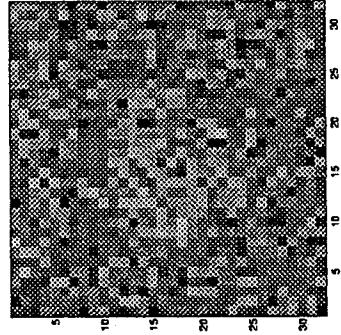
$$\frac{1}{T(K+1)} = \frac{1}{T(K)} + \Delta$$

Segmentation Example

- Iterated Conditional Modes (ICM): ML ; ICM 1; ICM 5; ICM 10



- Simulated Annealing (SA): ML ; SA 1; SA 5; SA 10



Other Bayesian Estimates

Y - observed image

X - segmentation

\hat{x} - estimate of segmentation

$C(X, \hat{x})$ = cost of choosing \hat{x} if X is the correct answer.

We would like to minimize risk in our choice.

$$E[C(X, \hat{x}) | Y=y] = R(y, \hat{x})$$

$$\hat{x} = \underset{x}{\operatorname{argmin}} R(y, x)$$

$$= \underset{x}{\operatorname{argmin}} E[C(X, x) | Y=y]$$

Choose

$$C_{\text{MAP}}(X, \hat{x}) = \begin{cases} 1 & X \neq \hat{x} \\ 0 & X = \hat{x} \end{cases}$$

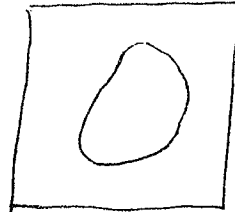
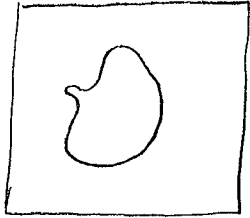
$$E[C_{\text{MAP}}(X, \hat{x}) | Y=y] = R(\hat{x}, y)$$

$$= 1 - P(X = \hat{x} | Y=y)$$

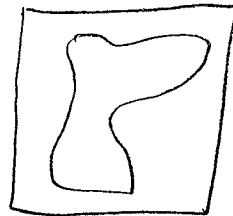
$$\hat{x} = \underset{x}{\operatorname{argmin}} \{1 - P(X=x | Y=y)\} \\ = \underset{x}{\operatorname{argmax}} P(X=x | Y=y) = \hat{x}_{\text{MAP}}$$

Problem with MAP estimate

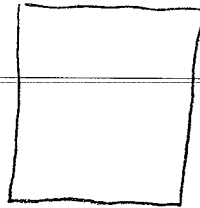
X



\hat{X}_a



\hat{X}_b



\hat{X}_c

All these have the same "cost"
But clearly \hat{X}_a is the best.

Lecture 22

Choose

$$C_{\text{MPM}}(X, \hat{x}) = \sum_{s \in S} t(X_s - \hat{x}_s)$$

$$\begin{aligned} R(\hat{x}, y) &= E[C_{\text{MPM}}(X, \hat{x}) \mid Y=y] \\ &= \sum_{s \in S} E[t(X_s - \hat{x}_s) \mid Y=y] \\ &= \sum_{s \in S} \{1 - P(X_s = \hat{x}_s \mid Y=y)\} \end{aligned}$$

$$\hat{x} = \underset{x}{\operatorname{argmax}} \sum_{s \in S} \{1 - P(X_s = x \mid Y=y)\}$$

$$= \underset{x}{\operatorname{argmin}} \sum_s P(X_s = x \mid Y=y)$$

$$\hat{x}_s = \underset{x_s}{\operatorname{argmin}} P(X_s = x_s \mid Y=y)$$

\hat{x}_s is the maximizer of the posterior marginal distribution (MPM)

How do we compute X_{MCM} ?

$$U(x) = \sum_{S \in \mathcal{S}} \ell(y | x_S) + \sum_{\{i, j\} \in \mathcal{C}} \beta \phi(x_i, x_j)$$

1. $p(x | y) = \frac{1}{Z} e^{-U(x)}$

1) $p(x_S | y) = \sum_{x_i, i \neq S} p(x_S, x_i, i \neq S, y)$
↑ impossible to compute

2) Generate $x^{(1)}, x^{(2)}, \dots$ from $p(x | y)$ using Gibbs sampler.

Then by ergodic property of Markov Chain

$$p(x_S | y) = \lim_{N \rightarrow \infty} \frac{1}{N} \underbrace{\sum_{k=0}^N \delta(x_S^{(k)} - x_S)}_{\# \text{ of times } x_S^{(k)} = x_S}$$

Pick

$$\hat{x}_S = \operatorname{argmax}_{x_S} p(x_S | y)$$

Problems with MPM

- 1) If β is large then N must be VERY large to achieve ergodic behavior. Computationally expensive
- 2) CPM requires no "consistency" in the solutions. Each pixel is chosen individually

Multi Scale Techniques

Idea

Make large scale "global" moves in search algorithms

- 1) Much faster computation
- 2) More sensitive to large scale patterns.

Define

$$t_1(X) = \sum_{\substack{i,j \\ i \neq j}} t(x_i, x_j)$$

= # of horizontal/vertical pairs of different values.

$$U(X) = \beta t_1(X)$$

- 1) Restrict X to have fixed value on 2×2 blocks

$$\begin{array}{cc|cc} 0 & 0 & 1 & 1 & & 0 & 1 \\ 0 & 0 & 1 & 1 & & & \\ \hline 1 & 1 & 0 & 0 & & 1 & 0 \\ 1 & 1 & 0 & 0 & & & \\ \hline & & X & & & & X^{(2)} \end{array}$$

- 2) Denote the new "decimated" field by $X^{(2)}$

$$U(X) = \beta t_1(X) = 2\beta t_1(X^{(2)})$$

$$\beta^{(2)} \stackrel{\Delta}{=} 2\beta$$

$$t_1 \stackrel{\Delta}{=} t_1(x)$$

$$t_2^{(2)} \stackrel{\Delta}{=} t_2(x^{(2)})$$

$$U(x) = \beta^{(2)} t_1^{(2)} = U^{(2)}(x^{(2)})$$

MRS algorithm

recursive \rightarrow

1. Use MRS algorithm to solve

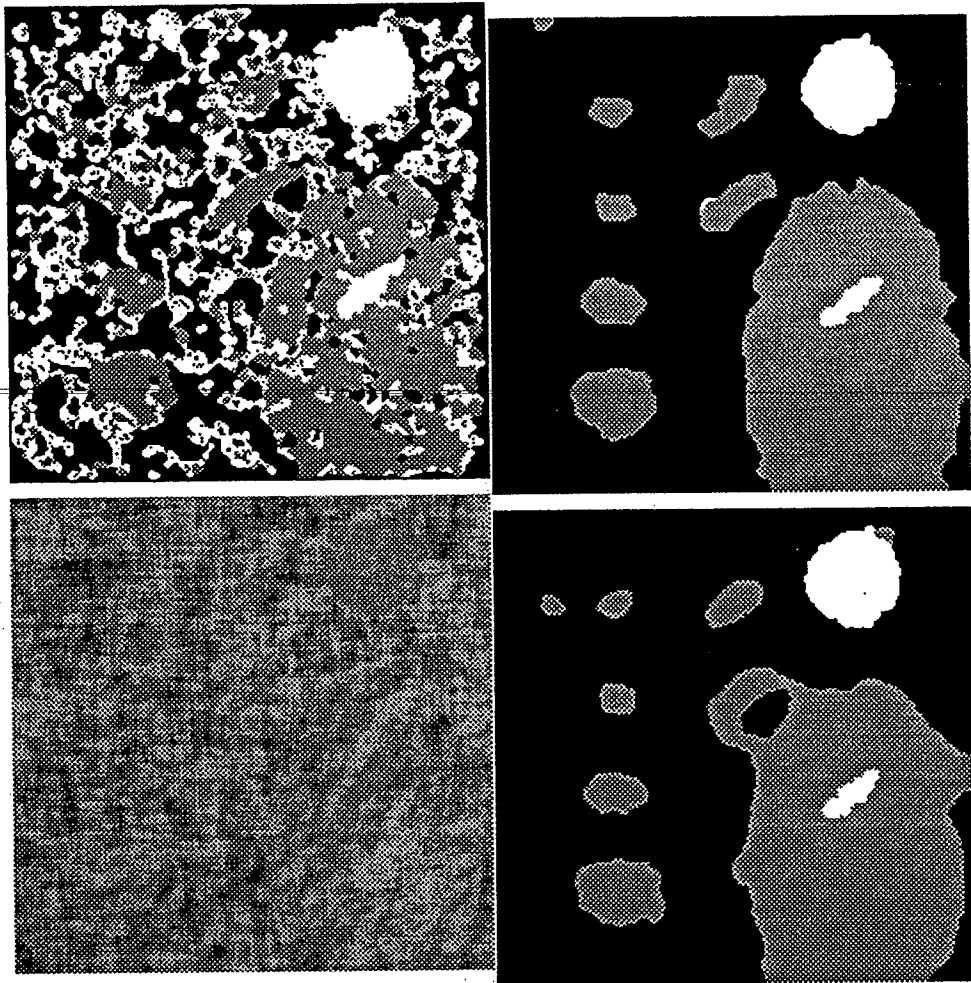
$$\hat{x}^{(2)} = \underset{x^{(2)}}{\operatorname{argmin}} U^{(2)}(x^{(2)})$$

2. Replicate pixels to form $x^{(1)}$ from $x^{(2)}$

3. Use ICM with the starting point $x^{(1)}$ to approximately solve

$$\hat{x}^{(1)} = \underset{x^{(1)}}{\operatorname{argmin}} U(x^{(1)})$$

Texture Segmentation Example



a	b
c	d

a) Synthetic image with 3 textures b) ICM - 29 iterations c) Simulated Annealing - 100 iterations d) Multiresolution - 7.8 iterations