

Simulation

- Topics to be covered:
 - Gibbs sampler
 - Metropolis sampler
 - Hastings-Metropolis sampler

Generating Samples from a Gibbs Distribution

- How do we generate a random variable X with a Gibbs distribution?

$$p(x) = \frac{1}{Z} \exp \{-U(x)\}$$

- Generally, this problem is difficult.
- Markov Chains can be generated sequentially
- Non-causal structure of MRF's makes simulation difficult.

Gibbs Sampler[4]

- Replace each point with a sample from its conditional distribution

$$p(x_s | x_i^k \ i \neq s) = p(x_s | x_{\partial s})$$

- Scan through all the points in the image.
- Advantage
 - Eliminates need for rejections \Rightarrow faster convergence
- Disadvantage
 - Generating samples from $p(x_s | x_{\partial s})$ can be difficult.

Gibbs Sampler Algorithm

Gibbs Sampler Algorithm:

1. Set $N = \#$ of pixels
2. Order the N pixels as $N = s(0), \dots, s(N - 1)$
3. Repeat for $k = 0$ to ∞
 - (a) Form $X^{(k+1)}$ from $X^{(k)}$ via

$$X_r^{(k+1)} = \begin{cases} W & \text{if } r = s(k) \\ X_r^{(k)} & \text{if } r \neq s(k) \end{cases}$$

where $W \sim p(x_{s(k)} | X_{\partial s(k)}^{(k)})$

The Metropolis Sampler[9, 8]

- How do we generate a sample from a Gibbs distribution?

$$p(x) = \frac{1}{Z} \exp \{-U(x)\}$$

- Start with the sample x^k , and generate a new sample W with probability $q(w|x^k)$.

Note: $q(w|x^k)$ must be symmetric.

$$q(w|x^k) = q(x^k|w)$$

- Compute $\Delta E(W) = U(W) - U(x^k)$, then do the following:

If $\Delta E(W) < 0$

– Accept: $X^{k+1} = W$

If $\Delta E(W) \geq 0$

– Accept: $X^{k+1} = W$ with probability $\exp\{-\Delta E(W)\}$

– Reject: $X^{k+1} = x^k$ with probability $1 - \exp\{-\Delta E(W)\}$

Ergodic Behavior of Metropolis Sampler

- The sequence of random fields, X^k , form a Markov chain.
- Let $p(x^{k+1}|x^k)$ be the transition probabilities of the Markov chain.
- Then X^k is reversible

$$p(x^{k+1}|x^k) \exp\{-U(x^k)\} = \exp\{-U(x^{k+1})\} p(x^k|x^{k+1})$$

- Therefore, if the Markov chain is irreducible, then

$$\lim_{k \rightarrow \infty} P\{X^k = x\} = \frac{1}{Z} \exp\{-U(x)\}$$

- If every state can be reached, then as $k \rightarrow \infty$, X^k will be a sample from the Gibbs distribution.

Example Metropolis Sampler for Ising Model

	0	
1	x_s	0
	0	

- Assume $x_s^k = 0$.
- Generate a binary R.V., W , such that $P\{W = 0\} = 0.5$.

$$\begin{aligned}\Delta E(W) &= U(W) - U(x_s^k) \\ &= \begin{cases} 0 & \text{if } W = 0 \\ 2\beta & \text{if } W = 1 \end{cases}\end{aligned}$$

If $\Delta E(W) < 0$

– Accept $X_s^{k+1} = W$

If $\Delta E(W) \geq 0$

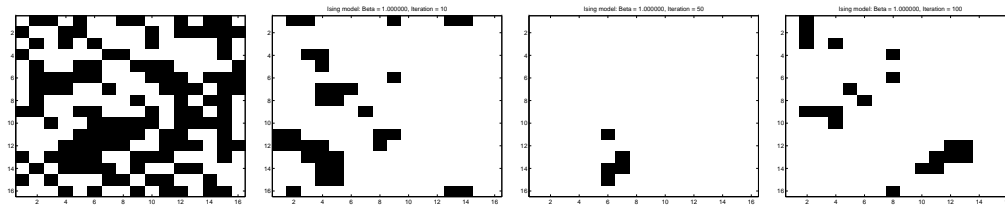
– Accept: $X_s^{k+1} = W$ with probability $\exp\{-\Delta E(W)\}$

– Reject: $X_s^{k+1} = x_s^k$ with probability $1 - \exp\{-\Delta E(W)\}$

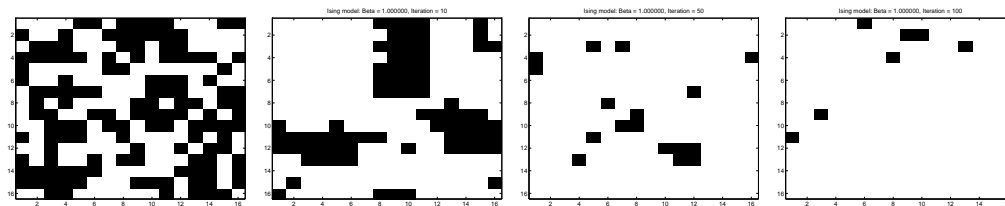
- Repeat this procedure for each pixel.
- **Warning:** for $\beta > \beta_c$ convergence can be extremely slow!

Example Simulation for Ising Model ($\beta = 1.0$)

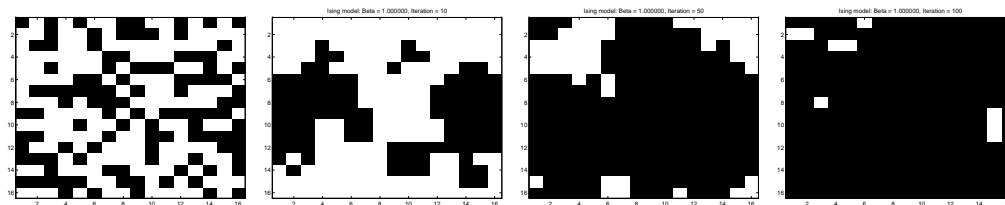
- Test 1



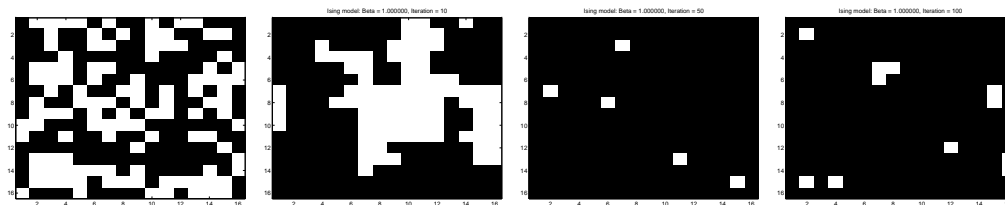
- Test 2



- Test 3



- Test 4



Uniform Random

10 Iterations

50 Iterations

100 Iterations

Advantages and Disadvantages of Metropolis Sampler

- Advantages
 - Can be implemented whenever ΔE is easy to compute.
 - Has guaranteed geometric convergence.
- Disadvantages
 - Can be slow if there are many rejections.
 - Is constrained to use a symmetric transition function $q(x^{k+1}|x^k)$.

Hastings-Metropolis Sampler[7, 10]

- Hastings and Peskun generalized the Metropolis sampler for transition functions $q(w|x^k)$ which are not symmetric.
- The acceptance probability is then

$$\alpha(x_s^k, w) = \min \left\{ 1, \frac{q(x^k|w)}{q(w|x^k)} \exp\{-\Delta E(w)\} \right\}$$

- Special cases

$$q(w|x^k) = q(x^k|z) \Rightarrow \text{conventional Metropolis}$$

$$q(w_s|x^k) = p(x_s^k|x_{\partial s}^k)|_{x_s^k=w_s} \Rightarrow \text{Gibbs sampler}$$

- Advantage

– Transition function may be chosen to minimize rejections[6]

Parameter Estimation for Discrete State MRF's

- Topics to be covered:
 - Why is it difficult?
 - Coding/maximum pseudolikelihood
 - Least squares

Why is Parameter Estimation Difficult?

- Consider the ML estimate of β for an Ising model.
- Remember that

$$t_1(x) = (\# \text{ horz. and vert. neighbors of different value.})$$

- Then the ML estimate of β is

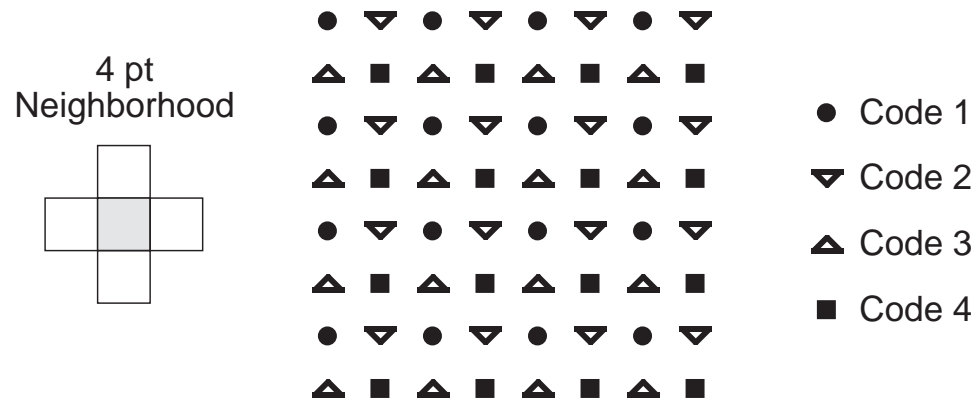
$$\begin{aligned}\hat{\beta} &= \arg \max_{\beta} \left\{ \frac{1}{Z(\beta)} \exp \{-\beta t_1(x)\} \right\} \\ &= \arg \max_{\beta} \{-\beta t_1(x) - \log Z(\beta)\}\end{aligned}$$

- However, $\log Z(\beta)$ has an intractable form

$$\log Z(\beta) = \log \sum_x \exp \{-\beta t_1(x)\}$$

- Partition function can not be computed.

Coding Method/Maximum Pseudolikelihood[1, 2]



- Assume a 4 point neighborhood
- Separate points into four groups or codes.
- Group (code) contains points which are conditionally independent given the other groups (codes).

$$\hat{\beta} = \arg \max_{\beta} \prod_{s \in \text{Code}_k} p(x_s | x_{\partial s})$$

- This is tractable (but not necessarily easy) to compute

Least Squares Parameter Estimation[3]

- It can be shown that for an Ising model

$$\log \frac{P\{X_s = 1|x_{\partial s}\}}{P\{X_s = 0|x_{\partial s}\}} = -\beta (V_1(1|x_{\partial s}) - V_1(0|x_{\partial s}))$$

- For each unique set of neighboring pixel values, $x_{\partial s}$, we may compute
 - The observed rate of $\log \frac{P\{X_s=1|x_{\partial s}\}}{P\{X_s=0|x_{\partial s}\}}$
 - The value of $(V_1(1|x_{\partial s}) - V_1(0|x_{\partial s}))$
 - This produces a set of over-determined linear equations which can be solved for β .
- This least squares method is easily implemented.

Theoretical Results in Parameter Estimation for MRF's

- Inconsistency of ML estimate for Ising model[11, 12]
 - Caused by critical temperature behavior.
 - Single sample of Ising model cannot distinguish between high β with mean $1/2$, and low β with large mean.
 - Not identifiable
- Consistency of maximum pseudolikelihood estimate[5]
 - Requires an identifiable parameterization.

References

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