

## Least Squares Adjustment with Rank-Deficient Weight Matrix

1. Model

$$\vec{y} = A\delta\vec{x} + \vec{d} + \vec{e} \quad \vec{e} \sim (0, \Sigma) \quad \text{where } \Sigma = \sigma_0^2 P^+ \text{ and } P\vec{d} = 0 \quad \text{A.1}$$

2. LSA Target Function  $\phi(\delta\vec{x}, \vec{d})$

$$\phi(\delta\vec{x}, \vec{d}) = \vec{e}^T P \vec{e} = (\vec{y} - A\delta\vec{x} - \vec{d})^T P (\vec{y} - A\delta\vec{x} - \vec{d}) = \min_{\delta\vec{x}, \vec{d}} \quad \text{A.2}$$

Since  $P\vec{d} = 0$ :

$$\phi(\delta\vec{x}, \vec{d}) = \phi(\delta\vec{x}) = \vec{e}^T P \vec{e} = (\vec{y} - A\delta\vec{x})^T P (\vec{y} - A\delta\vec{x}) = \min_{\delta\vec{x}} \quad \text{A.3}$$

Expanding Equation A.3 we get:

$$\begin{aligned} \phi(\delta\vec{x}) &= (\vec{y} - A\delta\vec{x})^T P (\vec{y} - A\delta\vec{x}) \\ &= \vec{y}^T P \vec{y} - \vec{y}^T P A \delta\vec{x} - \delta\vec{x}^T A^T P \vec{y} + \delta\vec{x}^T A^T P A \delta\vec{x} = \min_{\delta\vec{x}} \end{aligned} \quad \text{A.4}$$

Equation A.3 can be simplified to:

$$\phi(\delta\vec{x}) = \vec{y}^T P \vec{y} + \delta\vec{x}^T A^T P A \delta\vec{x} - 2\delta\vec{x}^T A^T P \vec{y} = \min_{\delta\vec{x}} \quad \text{A.5}$$

3. Solution Vector ( $\delta\hat{\vec{x}}$ )

The solution vector ( $\delta\hat{\vec{x}}$ ) that minimizes  $\phi(\delta\vec{x})$  can be obtained by differentiating  $\phi(\delta\vec{x})$  w.r.t.  $\delta\vec{x}$  and equating it to zero:

$$\frac{\partial \phi}{\partial \delta\vec{x}} = 2A^T P A \delta\vec{x} - 2A^T P \vec{y} = 0 \quad \text{A.6}$$

$$\delta\hat{\vec{x}} = (A^T P A)^{-1} A^T P \vec{y} = N^{-1} A^T P \vec{y} \quad \text{where } N = A^T P A \quad \text{A.7}$$

4. Variance-covariance matrix of the solution vector ( $\Sigma\{\delta\hat{\vec{x}}\}$ )

Using the law of error propagation, the variance-covariance matrix of the solution vector ( $\Sigma\{\delta\hat{\vec{x}}\}$ ) can be obtained as follows:

$$\Sigma\{\delta\hat{\vec{x}}\} = \sigma_0^2 N^{-1} A^T P^+ P^+ A N^{-1} \quad \text{A.8}$$

Since for a Moore-Penrose pseudo-inverse,  $P^+ P^+ P^+ = P^+$  (Koch, 1988):

$$\Sigma\{\delta\hat{\vec{x}}\} = \sigma_o^2 N^{-1} N N^{-1} = \sigma_o^2 N^{-1} \quad \text{A.9}$$

5. A-posteriori variance factor ( $\hat{\sigma}_o^2$ )

The a-posteriori variance factor ( $\hat{\sigma}_o^2$ ) is obtained by deriving the expected value of the sum of squares of the weighted predicted residuals:

$$E(\tilde{\vec{e}}^T P \tilde{\vec{e}}) = E\{(\vec{y} - A\widehat{\delta\vec{x}} - \vec{d})^T P (\vec{y} - A\widehat{\delta\vec{x}} - \vec{d})\} \quad \text{A.10}$$

Since  $P\vec{d} = 0$ , Equation A.10 gets the form:

$$E(\tilde{\vec{e}}^T P \tilde{\vec{e}}) = E\{(\vec{y} - A\widehat{\delta\vec{x}})^T P (\vec{y} - A\widehat{\delta\vec{x}})\} \quad \text{A.11}$$

Expanding Equation A.11 while using the derived solution for  $\widehat{\delta\vec{x}}$  in Equation A.7 we get (while considering that  $(I_n - AN^{-1}A^T P)$  is an idempotent matrix):

$$E(\tilde{\vec{e}}^T P \tilde{\vec{e}}) = E\{\vec{y}^T P \vec{y} - \vec{y}^T P AN^{-1}A^T P \vec{y}\} \quad \text{A.12}$$

Given that the trace of a scalar equals to the scalar, i.e.,  $tr(S) = S$  and that the trace operation is commutative, i.e.,  $tr(AB) = tr(BA)$  (Koch, 1988), Equation A.12 can be manipulated as follows:

$$E(\tilde{\vec{e}}^T P \tilde{\vec{e}}) = E\{tr(P \vec{y} \vec{y}^T) - tr(P AN^{-1}A^T P \vec{y} \vec{y}^T)\} \quad \text{A.13}$$

Based on the properties that  $tr(A) + tr(B) = tr(A+B)$  and that  $E\{tr(A)\} = trE(A)$  (Koch, 1988), Equation A.13 can be rewritten as follows:

$$\begin{aligned} E(\tilde{\vec{e}}^T P \tilde{\vec{e}}) &= trP [E(\vec{y} \vec{y}^T) - AN^{-1}A^T P E(\vec{y} \vec{y}^T)] \\ &= trP (I_n - AN^{-1}A^T P) E(\vec{y} \vec{y}^T) \end{aligned} \quad \text{A.14}$$

where:

$I_n$  is an  $n \times n$  identity matrix.

The term  $E(\vec{y} \vec{y}^T)$  can be derived from the variance-covariance matrix of the observations vector ( $\Sigma\{\vec{y}\}$ ) as follows:

$$\Sigma\{\vec{y}\} = \sigma_o^2 P^+ = E\{(\vec{y} - A\delta\vec{x} - \vec{d})(\vec{y} - A\delta\vec{x} - \vec{d})^T\} \quad \text{A.15}$$

Expanding Equation A.15, we get:

$$\begin{aligned} E(\vec{y}\vec{y}^T) &= \sigma_o^2 P^+ + (A \delta \vec{x} + \vec{d})(A \delta \vec{x} + \vec{d})^T \\ &= \sigma_o^2 P^+ + A \delta \vec{x} \delta \vec{x}^T A^T + A \delta \vec{x} \vec{d}^T + \vec{d} \delta \vec{x}^T A^T + \vec{d} \vec{d}^T \end{aligned} \quad \text{A.16}$$

Substituting Equation A.16 in Equation A.14 yields:

$$\begin{aligned} E(\tilde{\vec{e}}^T P \tilde{\vec{e}}) &= \text{tr} P (I - AN^{-1}A^T P) [\sigma_o^2 P^+ + A \delta \vec{x} \delta \vec{x}^T A^T + A \delta \vec{x} \vec{d}^T + \vec{d} \delta \vec{x}^T A^T \\ &\quad + \vec{d} \vec{d}^T] \end{aligned} \quad \text{A.17}$$

Given that  $P \vec{d} = 0$  and  $(I - AN^{-1}A^T P)A = 0$ , Equation A.17 can be simplified to:

$$E(\tilde{\vec{e}}^T P \tilde{\vec{e}}) = \sigma_o^2 \text{tr} P (I - AN^{-1}A^T P) P^+ = \sigma_o^2 \text{tr} P P^+ - \sigma_o^2 \text{tr} N^{-1} A^T P P^+ P A \quad \text{A.18}$$

Based on the property that  $\text{tr}(AB) = \text{rank}(AB)$  (given that  $AB$  is idempotent) and  $\text{rank}(AB) \leq \min(\text{rank}A, \text{rank}B)$  (Koch, 1988), the following can be stated:

$$\text{tr}(P P^+) = \text{rank}(P P^+) = \min(\text{rank}P, \text{rank}P^+) = \text{rank}P = q \quad \text{A.19}$$

Given that  $\text{tr}(P P^+) = q$  (as shown in Equation A.19) and that  $P P^+ P = P$ , Equation A.18 can be simplified to:

$$E(\tilde{\vec{e}}^T P \tilde{\vec{e}}) = \sigma_o^2 q - \sigma_o^2 \text{tr} N^{-1} N = \sigma_o^2 q - \sigma_o^2 \text{tr} I_m = \sigma_o^2 q - \sigma_o^2 m \quad \text{A.20}$$

where,

$m$  is the number of unknown parameters.

Finally, we can get the expression for the estimated a-posteriori variance factor ( $\hat{\sigma}_o^2$ ) from Equation A.20 as follows:

$$\hat{\sigma}_o^2 = \frac{\tilde{\vec{e}}^T P \tilde{\vec{e}}}{(\text{rank}P - m)} \quad \text{A.21}$$